

A Derivations of Variance Controlled Diffusion SDEs

A.1 Proof of Proposition 4.1

Proposition 4.1 For any bounded measurable function $\tau(t) : [0, T] \rightarrow \mathbb{R}$, the following Reverse SDEs

$$d\mathbf{x}_t = \left[f(t)\mathbf{x}_t - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t) \nabla_{\mathbf{x}} \log p_t(\mathbf{x}_t) \right] dt + \tau(t)g(t)d\bar{\mathbf{w}}_t, \quad \mathbf{x}_T \sim p_T(\mathbf{x}_T) \quad (20)$$

has the same marginal probability distributions with (4) and (3).

Proof. Denote $p(\mathbf{x}, t) : \mathbb{R}^d \times [0, T] \rightarrow \mathbb{R}^+$ as the probability density function of \mathbf{x}_t at time t , thus $p(\mathbf{x}, T) = p_T(\mathbf{x})$. Fokker-Planck equation [38] determines a Partial Differential Equation (PDE) that $p_t(\mathbf{x})$ satisfies:

$$\begin{aligned} -\frac{\partial p_t(\mathbf{x})}{\partial t} &= -\sum_{i=1}^d \frac{\partial}{\partial x_i} \left[-\left[f(t)p_t(\mathbf{x})x_i - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right] \right] \\ &\quad + \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d \frac{\partial^2}{\partial x_i \partial x_j} [\tau^2(t)g^2(t)\delta_{ij}p_t(\mathbf{x})] \\ &= \sum_{i=1}^d \frac{\partial}{\partial x_i} \left[f(t)p_t(\mathbf{x})x_i - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right] \\ &\quad + \frac{1}{2} \sum_{i=1}^d \frac{\partial}{\partial x_i} \left[\frac{\partial}{\partial x_i} [\tau^2(t)g^2(t)p_t(\mathbf{x})] \right]. \end{aligned} \quad (21)$$

Eq. (20) is a reverse-time SDE running from T to 0, thus there are two additional minus signs in Eq. (21) before term $\frac{\partial p_t(\mathbf{x})}{\partial t}$ and term $\left[f(t)p_t(\mathbf{x})x_i - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right]$ compared with vanilla Fokker-Planck equation in general cases. Here δ_{ij} is the Dirac symbol satisfies $\delta_{ij} = 1$ when $i = j$, otherwise, $\delta_{ij} = 0$. Notice that

$$\frac{\partial}{\partial x_i} [\tau^2(t)g^2(t)p_t(\mathbf{x})] = \tau^2(t)g^2(t) \frac{\partial}{\partial x_i} p_t(\mathbf{x}) = \tau^2(t)g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i}. \quad (22)$$

Substituting Eq. (22) into Eq. (21), we obtain that

$$\begin{aligned} -\frac{\partial p_t(\mathbf{x})}{\partial t} &= \sum_{i=1}^d \frac{\partial}{\partial x_i} \left[f(t)p_t(\mathbf{x})x_i - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right] \\ &\quad + \frac{1}{2} \sum_{i=1}^d \frac{\partial}{\partial x_i} \left[\tau^2(t)g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right] \\ &= \sum_{i=1}^d \frac{\partial}{\partial x_i} \left[f(t)p_t(\mathbf{x})x_i - \frac{1}{2}g^2(t)p_t(\mathbf{x}) \frac{\partial \log p_t(\mathbf{x})}{\partial x_i} \right], \end{aligned} \quad (23)$$

which is independent of $\tau(t)$. With the same initial condition $p(\mathbf{x}, T) = p_T(\mathbf{x})$, the family of Reverse SDEs in Eq. (20) have exactly the same evolutions of probability density function because they share the same Fokker-Planck equation. Especially, when $\tau(t) = 0$, Eq. (20) degenerates to diffusion ODEs and when $\tau(t) = 1$, Eq. (20) degenerates to diffusion SDEs. \square

A.2 Two Reparameterizations and Exact Solution under Exponential Integrator

In this subsection, we will show the exact solution of SDE in both *data prediction* reparameterization and *noise prediction* reparameterization. The noise term in *data prediction* has smaller variance than *noise prediction* ones, implying the necessity of adopting *data prediction* reparameterization for the SDE sampler.

A.2.1 Data Prediction Reparameterization

After approximating $\nabla_{\mathbf{x}} \log p_t(\mathbf{x}_t)$ with $\mathbf{s}_\theta(\mathbf{x}_t, t)$ and reparameterizing $\mathbf{s}_\theta(\mathbf{x}_t, t)$ with $-(\mathbf{x}_t - \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t))/\sigma_t^2$, Eq. (20) becomes

$$d\mathbf{x}_t = \left[f(t)\mathbf{x}_t + \left(\frac{1 + \tau^2(t)}{2\sigma_t} \right) g^2(t) \left(\frac{\mathbf{x}_t - \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t)}{\sigma_t} \right) \right] dt + \tau(t)g(t)d\bar{\mathbf{w}}_t. \quad (24)$$

Applying change-of-variable with log-SNR $\lambda_t = \log(\alpha_t/\sigma_t)$ and substituting the following relationship

$$f(t) = \frac{d \log \alpha_t}{dt}, \quad g^2(t) = \frac{d\sigma_t^2}{dt} - 2 \frac{d \log \alpha_t}{dt} \sigma_t^2 = -2\sigma_t^2 \frac{d\lambda_t}{dt}, \quad (25)$$

Eq. (24) becomes

$$\begin{aligned} d\mathbf{x}_t &= \left[\frac{d \log \alpha_t}{dt} \mathbf{x}_t - (1 + \tau^2(t)) (\mathbf{x}_t - \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t)) \frac{d\lambda_t}{dt} \right] dt + \tau(t)\sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t \\ &= \left[\left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) \mathbf{x}_t + (1 + \tau^2(t)) \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} \right] dt \\ &\quad + \tau(t)\sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t. \end{aligned} \quad (26)$$

A.2.2 Proof of Proposition 4.2

Proposition 4.2 Given \mathbf{x}_s for any time $s > 0$, the solution \mathbf{x}_t at time $t \in [0, s]$ of Eq. (9) is

$$\begin{aligned} \mathbf{x}_t &= \frac{\sigma_t}{\sigma_s} e^{-\int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \mathbf{x}_s + \sigma_t \mathbf{F}_\theta(s, t) + \sigma_t \mathbf{G}(s, t), \\ \mathbf{F}_\theta(s, t) &= \int_{\lambda_s}^{\lambda_t} e^{-\int_{\lambda}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} (1 + \tau^2(\lambda)) e^\lambda \mathbf{x}_\theta(\mathbf{x}_\lambda, \lambda) d\lambda \\ \mathbf{G}(s, t) &= \int_s^t e^{-\int_{\lambda_u}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u, \end{aligned} \quad (27)$$

where $\mathbf{G}(s, t)$ is an Itô integral [38] with the special property

$$\sigma_t \mathbf{G}(s, t) \sim \mathcal{N}\left(\mathbf{0}, \sigma_t^2 \left(1 - e^{-2 \int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}}\right)\right). \quad (28)$$

Proof. Define $\mathbf{y}_t = e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \mathbf{x}_t$, where $t_0 \in [0, T]$ is a constant. Differentiate \mathbf{y}_t with respect to t , we get

$$\begin{aligned} d\mathbf{y}_t &= - \left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \mathbf{x}_t dt \\ &\quad + e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} d\mathbf{x}_t \\ &= - \left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \mathbf{x}_t dt \\ &\quad + e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \left[\left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) \mathbf{x}_t \right] dt \\ &\quad + e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \left[(1 + \tau^2(t)) \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} \right] dt \\ &\quad + e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t \\ &= e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \left[(1 + \tau^2(t)) \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} \right] dt \\ &\quad + e^{-\int_{t_0}^t \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv}\right) dv} \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t. \end{aligned} \quad (29)$$

Integrating both sides from s to t

$$\begin{aligned} \mathbf{y}_t = \mathbf{y}_s &+ \int_s^t e^{-\int_{t_0}^u \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv} \right) dv} \left[(1 + \tau^2(u)) \alpha_u \mathbf{x}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} \right] du \\ &+ \int_s^t e^{-\int_{t_0}^u \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv} \right) dv} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \end{aligned} \quad (30)$$

Substituting the definition of $\mathbf{y}_t, \mathbf{y}_s$ into Eq. (30), we obtain Eq. (27)

$$\begin{aligned} \mathbf{x}_t &= e^{\int_s^t \left(\frac{d \log \alpha_u}{du} - (1 + \tau^2(u)) \frac{d\lambda_u}{du} \right) du} \mathbf{x}_s \\ &+ \int_s^t e^{-\int_t^u \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv} \right) dv} \left[(1 + \tau^2(u)) \alpha_u \mathbf{x}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} \right] du \\ &+ \int_s^t e^{-\int_t^u \left(\frac{d \log \alpha_v}{dv} - (1 + \tau^2(v)) \frac{d\lambda_v}{dv} \right) dv} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \\ &= e^{\int_s^t \left(\frac{d \log \sigma_u}{du} - \tau^2(u) \frac{d\lambda_u}{du} \right) du} \mathbf{x}_s \\ &+ \int_s^t e^{-\int_t^u \left(\frac{d \log \sigma_v}{dv} - \tau^2(v) \frac{d\lambda_v}{dv} \right) dv} \left[(1 + \tau^2(u)) \alpha_u \mathbf{x}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} \right] du \\ &+ \int_s^t e^{-\int_t^u \left(\frac{d \log \sigma_v}{dv} - \tau^2(v) \frac{d\lambda_v}{dv} \right) dv} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \\ &= \frac{\sigma_t}{\sigma_s} e^{-\int_s^t \tau^2(u) \frac{d\lambda_u}{du}} \mathbf{x}_s \\ &+ \int_s^t \frac{\sigma_t}{\sigma_u} e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv}} \left[(1 + \tau^2(u)) \alpha_u \mathbf{x}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} \right] du \\ &+ \int_s^t \frac{\sigma_t}{\sigma_u} e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv}} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \\ &= \frac{\sigma_t}{\sigma_s} e^{-\int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \mathbf{x}_s \\ &+ \sigma_t \int_{\lambda_s}^{\lambda_t} e^{-\int_{\bar{\lambda}}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} (1 + \tau^2(\lambda)) e^\lambda \mathbf{x}_\theta(\mathbf{x}_\lambda, \lambda) d\lambda \\ &+ \sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv}} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \end{aligned} \quad (31)$$

The last term $\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv}} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u$ of Eq. (31) is the $It\hat{o}$ integral term. It follows a Gaussian distribution, which can be directly derived from two basic facts [38]: first, the definition of $It\hat{o}$ integral is the limitation in L^2 space; second, the limit of Gaussian Process in L^2 space is still Gaussian. Then we can compute the mean as:

$$\mathbb{E} \left[\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv}} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] = 0 \quad (32)$$

and the variance is

$$\begin{aligned}
& \text{Var} \left[\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] \\
&= \mathbb{E} \left[\left(\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right)^2 \right] - \left(\mathbb{E} \left[\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] \right)^2 \\
&= \mathbb{E} \left[\left(\sigma_t \int_t^s e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\mathbf{w}_u \right)^2 \right] - 0 \\
&= \sigma_t^2 \mathbb{E} \left[\int_t^s \left(e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} \right)^2 du \right] \\
&= \sigma_t^2 \int_t^s e^{-\int_u^t 2\tau^2(v) \frac{d\lambda_v}{dv} dv} \tau^2(u) \left(-2 \frac{d\lambda_u}{du} \right) du \\
&= \sigma_t^2 \int_{\lambda_s}^{\lambda_t} 2e^{-\int_{\lambda}^{\lambda_t} 2\tau^2(\bar{\lambda}) d\bar{\lambda}} \tau^2(\lambda) d\lambda
\end{aligned} \tag{33}$$

The expectation equals zero because the Itô integral is a martingale [38]. The computation of variance uses the *Itô Isometry*, which is a crucial fact of Itô integral. We can further simplify the result by using the change of variable $P(\lambda) = e^{\int_{\lambda}^{\lambda_t} 2\tau^2(\bar{\lambda}) d\bar{\lambda}}$.

$$\begin{aligned}
& \text{Var} \left[\sigma_t \int_s^t e^{-\int_u^t \tau^2(v) \frac{d\lambda_v}{dv} dv} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] \\
&= \sigma_t^2 \int_{\lambda_s}^{\lambda_t} 2e^{-\int_{\lambda}^{\lambda_t} 2\tau^2(\bar{\lambda}) d\bar{\lambda}} \tau^2(\lambda) d\lambda \\
&= \sigma_t^2 \int_{P(\lambda_s)}^{P(\lambda_t)} P(\lambda) \frac{dP(\lambda)}{P(\lambda)} \\
&= \sigma_t^2 (P(\lambda_t) - P(\lambda_s)) \\
&= \sigma_t^2 (1 - e^{-2 \int_{\lambda_s}^{\lambda_t} \tau^2(\lambda) d\lambda})
\end{aligned} \tag{34}$$

□

A.2.3 Noise Prediction Reparameterization

After approximating $\nabla_{\mathbf{x}} \log p_t(\mathbf{x}_t)$ with $\mathbf{s}_{\theta}(\mathbf{x}_t, t)$ and reparameterizing $\mathbf{s}_{\theta}(\mathbf{x}_t, t)$ with $-\epsilon_{\theta}(\mathbf{x}_t, t)/\sigma_t$, Eq. (20) becomes

$$d\mathbf{x}_t = \left[f(t)\mathbf{x}_t + \left(\frac{1 + \tau^2(t)}{2\sigma_t} \right) g^2(t)\epsilon_{\theta}(\mathbf{x}_t, t) \right] dt + \tau(t)g(t)d\bar{\mathbf{w}}_t. \tag{35}$$

Applying change-of-variable with log-SNR $\lambda_t = \log(\alpha_t/\sigma_t)$ and substituting the following relationship

$$f(t) = \frac{d \log \alpha_t}{dt}, \quad g^2(t) = \frac{d\sigma_t^2}{dt} - 2 \frac{d \log \alpha_t}{dt} \sigma_t^2 = -2\sigma_t^2 \frac{d\lambda_t}{dt}, \tag{36}$$

Eq. (35) becomes

$$d\mathbf{x}_t = \left[\frac{d \log \alpha_t}{dt} \mathbf{x}_t - (1 + \tau^2(t)) \sigma_t \epsilon_{\theta}(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} \right] dt + \tau(t)\sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t. \tag{37}$$

Eq (35) is the formulation of *noist prediction model*. Similar with Proposition 4.2, Eq. (37) can be solved analytically, which is shown in the following proposition:

Proposition A.1. Given \mathbf{x}_s for any time $s > 0$, the solution \mathbf{x}_t at time $t \in [0, s]$ of (37) is

$$\begin{aligned}\mathbf{x}_t &= \frac{\alpha_t}{\alpha_s} \mathbf{x}_s + \alpha_t \mathbf{F}_\theta(s, t) + \alpha_t \mathbf{G}(s, t), \\ \mathbf{F}_\theta(s, t) &= \int_{\lambda_s}^{\lambda_t} e^{-\lambda} (1 + \tau^2(\lambda)) \boldsymbol{\epsilon}_\theta(\mathbf{x}_\lambda, \lambda) d\lambda \\ \mathbf{G}(s, t) &= \int_s^t e^{-\lambda_u} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u,\end{aligned}\tag{38}$$

where $\mathbf{G}(s, t)$ is an Itô integral [38] with the special property

$$\alpha_t \mathbf{G}(s, t) \sim \mathcal{N}\left(\mathbf{0}, \alpha_t^2 \int_{\lambda_s}^{\lambda_t} 2e^{-2\lambda} \tau^2(\lambda) d\lambda\right).\tag{39}$$

Proof. Define $\mathbf{y}_t = e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \mathbf{x}_t$, where $t_0 \in [0, T]$ is a constant. Differentiate \mathbf{y}_t with respect to t , we get

$$\begin{aligned}d\mathbf{y}_t &= -\frac{d \log \alpha_t}{dt} e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \mathbf{x}_t dt + e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} d\mathbf{x}_t \\ &= -\frac{d \log \alpha_t}{dt} e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \mathbf{x}_t dt + e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \frac{d \log \alpha_t}{dt} \mathbf{x}_t dt \\ &\quad - e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} (1 + \tau^2(t)) \sigma_t \boldsymbol{\epsilon}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} dt + e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t \\ &= -e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} (1 + \tau^2(t)) \sigma_t \boldsymbol{\epsilon}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} dt + e^{-\int_{t_0}^t \frac{d \log \alpha_v}{dv} dv} \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t.\end{aligned}\tag{40}$$

Integrating both sides from s to t

$$\begin{aligned}\mathbf{y}_t &= \mathbf{y}_s - \int_s^t e^{-\int_{t_0}^u \frac{d \log \alpha_v}{dv} dv} (1 + \tau^2(u)) \sigma_u \boldsymbol{\epsilon}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} du \\ &\quad + \int_s^t e^{-\int_{t_0}^u \frac{d \log \alpha_v}{dv} dv} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u.\end{aligned}\tag{41}$$

Substituting the definition of $\mathbf{y}_t, \mathbf{y}_s$ into Eq. (41), we obtain

$$\begin{aligned}\mathbf{x}_t &= e^{\int_s^t \frac{d \log \alpha_u}{du} du} \mathbf{x}_s + \int_s^t e^{-\int_t^u \frac{d \log \alpha_v}{dv} dv} (1 + \tau^2(u)) \sigma_u \boldsymbol{\epsilon}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} du \\ &\quad + \int_s^t e^{-\int_t^u \frac{d \log \alpha_v}{dv} dv} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \\ &= \frac{\alpha_t}{\alpha_s} \mathbf{x}_s + \int_s^t \frac{\alpha_t}{\alpha_u} (1 + \tau^2(u)) \sigma_u \boldsymbol{\epsilon}_\theta(\mathbf{x}_u, u) \frac{d\lambda_u}{du} du + \int_s^t \frac{\alpha_t}{\alpha_u} \tau(u) \sigma_u \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \\ &= \frac{\alpha_t}{\alpha_s} \mathbf{x}_s + \alpha_t \int_{\lambda_s}^{\lambda_t} e^{-\lambda} (1 + \tau^2(\lambda)) \boldsymbol{\epsilon}_\theta(\mathbf{x}_\lambda, \lambda) d\lambda + \alpha_t \int_s^t e^{-\lambda_u} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u.\end{aligned}\tag{42}$$

The Itô integral term $\alpha_t \int_s^t e^{-\lambda_u} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u$ follows a Gaussian distribution. Following the derivation in Proposition 4.2, the mean of the Itô integral term is:

$$\mathbb{E} \left[\alpha_t \int_s^t e^{-\lambda_u} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] = 0\tag{43}$$

and the expectation is

$$\begin{aligned}
& \text{Var} \left[\alpha_t \int_s^t e^{-\lambda_u \tau(u)} \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] \\
&= \mathbb{E} \left[\left(\alpha_t \int_s^t e^{-\lambda_u \tau(u)} \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right)^2 \right] - \left(\mathbb{E} \left[\alpha_t \int_s^t e^{-\lambda_u \tau(u)} \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \right] \right)^2 \\
&= \mathbb{E} \left[\left(\alpha_t \int_t^s e^{-\lambda_u \tau(u)} \sqrt{-2 \frac{d\lambda_u}{du}} d\mathbf{w}_u \right)^2 \right] - 0 \\
&= \alpha_t^2 \mathbb{E} \left[\int_t^s \left(e^{-\lambda_u \tau(u)} \sqrt{-2 \frac{d\lambda_u}{du}} \right)^2 du \right] \\
&= \alpha_t^2 \int_t^s e^{-2\lambda_u \tau^2(u)} du \left(-2 \frac{d\lambda_u}{du} \right) du \\
&= \alpha_t^2 \int_{\lambda_s}^{\lambda_t} 2e^{-2\lambda \tau^2(\lambda)} d\lambda
\end{aligned} \tag{44}$$

□

A.2.4 Comparison between Data and Noise Reparameterizations

In Table 1 we perform an ablation study on data and noise reparameterizations, the experiment results show that under the same magnitude of stochasticity, the proposed *SA-Solver* in data reparameterization has a better convergence which leads to better FID results under the same NFEs. In this subsection, we provide a theoretical view of this phenomenon.

Corollary A.2. *For any bounded measurable function $\tau(t)$, the following inequality holds*

$$\sigma_t^2 \left(1 - e^{-2 \int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \right) \leq \alpha_t^2 \int_{\lambda_s}^{\lambda_t} 2e^{-2\lambda \tau^2(\lambda)} d\lambda. \tag{45}$$

Proof. It's equivalent to show that

$$1 - e^{-2 \int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \leq e^{2\lambda_t} \int_{\lambda_s}^{\lambda_t} 2e^{-2\lambda \tau^2(\lambda)} d\lambda. \tag{46}$$

From the basic inequality $1 - e^{-x} \leq x$, we have

$$1 - e^{-2 \int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \leq 2 \int_{\lambda_s}^{\lambda_t} \tau^2(\lambda) d\lambda. \tag{47}$$

Thus it's sufficient to show that

$$e^{2\lambda_t} \int_{\lambda_s}^{\lambda_t} 2e^{-2\lambda \tau^2(\lambda)} d\lambda \geq 2 \int_{\lambda_s}^{\lambda_t} \tau^2(\lambda) d\lambda, \tag{48}$$

which is true because

$$\int_{\lambda_s}^{\lambda_t} 2 \left(e^{2(\lambda_t - \lambda)} - 1 \right) \tau^2(\lambda) d\lambda \geq 0, \tag{49}$$

□

This corollary indicates that the same SDE under two different reparameterizations has different properties under the effect of the exponential integrator. Specifically, in the numerical scheme, the data reparameterization will inject smaller noise in each step's updation. We speculate that this is the reason that the data reparameterization has a better convergence, shown as in Table 1.

B Derivations and Proofs for SA-Solver

B.1 Preliminary

We will first review some basic concepts and formulas in the numerical solutions of SDEs [39]. Suppose we have an *Itô* SDE $d\mathbf{x}_t = f(\mathbf{x}_t, t)dt + g(\mathbf{x}_t, t)d\mathbf{w}_t$ and time steps $\{t_i\}_{i=0}^M, t_i \in [0, T]$ to numerically solve the SDE. For a random variable Z , we define the L_1 norm $\|Z\|_{L_1} = \mathbb{E}[|Z|]$, the L_2 norm $\|Z\|_{L_2} = \mathbb{E}[|Z|^2]^{\frac{1}{2}}$, where $|\cdot|$ is the Euclidean norm. Denote $h = \max_{1 \leq i \leq M} (t_i - t_{i-1})$.

Definition B.1. We shall say that a time-discrete approximation $\mathbf{x}_0, \dots, \mathbf{x}_M$, where \mathbf{x}_i is a numerical approximation of \mathbf{x}_{t_i} , converges strongly with order $\gamma > 0$, if there exists a positive constant C , which does not depend on h and a $h_0 > 0$ such that

$$\max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_1} \leq Ch^\gamma, \quad \forall h \leq h_0. \quad (50)$$

Definition B.2. We say it is mean-square convergent with order $\gamma > 0$, if there exists a positive constant C , which does not depend on h and a $h_0 > 0$ such that

$$\max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_2} \leq Ch^\gamma, \quad \forall h \leq h_0. \quad (51)$$

Remark 2. To prove the strong convergence order γ of a numerical scheme, it's sufficient to show the mean-square convergence order γ . This is from Hölder Inequality $\mathbb{E}[|Z|] \leq \mathbb{E}[|Z|^2]^{\frac{1}{2}} \mathbb{E}[|1|^2]^{\frac{1}{2}} \leq \mathbb{E}[|Z|^2]^{\frac{1}{2}}$. Thus $\max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_1} \leq \max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_2}$.

We also need the following definition and assumptions, which usually holds in practical diffusion models.

Definition B.3. A function $h : \mathbb{R}^d \times [0, T] \rightarrow \mathbb{R}^d$ satisfies a linear growth condition if there exists a constant K such that

$$|h(\mathbf{x}, t)| \leq K(1 + |\mathbf{x}|^2)^{\frac{1}{2}} \quad (52)$$

Assumption B.4. The data prediction model \mathbf{x}_θ and its derivatives such as $\partial_t \mathbf{x}_\theta, \nabla_{\mathbf{x}} \mathbf{x}_\theta$ and $\Delta \mathbf{x}_\theta$ satisfy the linear growth condition.

Assumption B.5. The data prediction model \mathbf{x}_θ satisfies a uniform Lipschitz condition with respect to \mathbf{x}

$$|\mathbf{x}_\theta(\mathbf{x}_1, t) - \mathbf{x}_\theta(\mathbf{x}_2, t)| \leq L|\mathbf{x}_1 - \mathbf{x}_2|, \quad \forall \mathbf{x}, \mathbf{y} \in \mathbb{R}^d, t \in [0, T] \quad (53)$$

B.2 Outline of the Proof

In the remaining part of this section, we will focus on our variance controlled SDE

$$d\mathbf{x}_t = \left[\left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) \mathbf{x}_t + (1 + \tau^2(t)) \alpha_t \mathbf{x}_\theta(\mathbf{x}_t, t) \frac{d\lambda_t}{dt} \right] dt + \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} d\bar{\mathbf{w}}_t. \quad (54)$$

Consider the general case of the numerical scheme as follows:

$$\begin{aligned} \mathbf{x}_{i+1} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda_u) d\lambda_u} \mathbf{x}_i + \sum_{j=-1}^{s-1} b_{i-j} \mathbf{x}_\theta(\mathbf{x}_{i-j}, t_{i-j}) \\ &\quad + \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_u^{t_{i+1}} \tau^2(\lambda) d\lambda} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \end{aligned} \quad (55)$$

in which Eq. (17) and Eq (14) are the special case of this scheme. We will provide proof of the mean-square convergence order of the numerical scheme $\max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_2}$. We define the local

error of the numerical scheme Eq. (55) for the approximation of the SDE Eq. (54) as

$$\begin{aligned} L_{i+1} = \mathbf{x}_{t_{i+1}} - \mathbf{x}_{i+1} &= \mathbf{x}_{t_{i+1}} - \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} \mathbf{x}_{t_i} - \sum_{j=-1}^{s-1} b_{i-j} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i-j}}, t_{i-j}) \\ &\quad - \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_u^{t_{i+1}} \tau^2(\lambda) d\lambda} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u. \end{aligned} \quad (56)$$

L_{i+1} can be decomposed into R_{i+1} and S_{i+1} . Then the mean-square convergence can be derived, which is summarized in the following theorem proved by [29]:

Theorem B.6 ([29], Theorem 1). *The mean-square convergent of \mathbf{x}_i is bounded by*

$$\max_{0 \leq i \leq M} \|\mathbf{x}_{t_i} - \mathbf{x}_i\|_{L_2} \leq S \left\{ \max_{0 \leq i \leq s-1} \|D_i\|_{L_2} + \max_{s \leq i \leq M} \left(\frac{\|R_i\|_{L_2}}{h} + \frac{\|S_i\|_{L_2}}{h^{\frac{1}{2}}} \right) \right\}. \quad (57)$$

In Eq. (57), $D_i, i = 0, \dots, s-1$ are the initial error which we do not consider. Given Theorem B.6, to show the convergence order $\mathcal{O}(\max_{0 \leq t \leq T} \tau(t)h + h^s)$ of our s -step *SA-Predictor* and the convergence order $\mathcal{O}(\max_{0 \leq t \leq T} \tau(t)h + h^{s+1})$ of our s -step *SA-Corrector*, we just need to prove the following lemmas.

Lemma B.7 (Convergence rate of s -step *SA-Predictor*). *For*

$$\begin{aligned} \mathbf{x}_{t_{i+1}} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda_u) d\lambda_u} \mathbf{x}_{t_i} + \sum_{j=0}^{s-1} b_{i-j} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i-j}}, t_{i-j}) + \tilde{\sigma}_i \boldsymbol{\xi}, \quad \boldsymbol{\xi} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}), \\ \tilde{\sigma}_i &= \sigma_{t_{i+1}} \sqrt{1 - e^{-2 \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda}} \\ b_{i-j} &= \sigma_{t_{i+1}} \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda_v) d\lambda_v} (1 + \tau^2(\lambda_u)) e^{\lambda_u} l_{i-j}(\lambda_u) d\lambda_u, \quad \forall 0 \leq j \leq s-1 \end{aligned} \quad (58)$$

There exists an decomposition of local error L_i such that $L_i = R_i + S_i$ and

$$\|R_i\|_{L_2} \leq h^{s+1}, \|S_i\|_{L_2} \leq \max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}, \quad (59)$$

Lemma B.8 (Convergence rate of s -step *SA-Corrector*). *For*

$$\begin{aligned} \mathbf{x}_{t_{i+1}} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda_u) d\lambda_u} \mathbf{x}_{t_i} + \hat{b}_{i+1} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i+1}}^p, t_{i+1}) + \sum_{j=0}^{s-1} \hat{b}_{i-j} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i-j}}, t_{i-j}) + \tilde{\sigma}_i \boldsymbol{\xi}, \\ \tilde{\sigma}_i &= \sigma_{t_{i+1}} \sqrt{1 - e^{-2 \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda}} \\ \hat{b}_{i-j} &= \sigma_{t_{i+1}} \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda_v) d\lambda_v} (1 + \tau^2(\lambda_u)) e^{\lambda_u} \hat{l}_{i-j}(\lambda_u) d\lambda_u, \quad \forall 0 \leq j \leq s-1 \\ \hat{b}_{i+1} &= \sigma_{t_{i+1}} \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda_v) d\lambda_v} (1 + \tau^2(\lambda_u)) e^{\lambda_u} \hat{l}_{i+1}(\lambda_u) d\lambda_u \end{aligned} \quad (60)$$

There exists an decomposition of local error L_i such that $L_i = R_i + S_i$ and

$$\|R_i\|_{L_2} \leq h^{s+2}, \|S_i\|_{L_2} \leq \max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}, \quad (61)$$

Lemma B.7 and B.8 will be proved in Sec. B.4.

B.3 Lemmas for the Proof

To better analyze the local error here, we state the following definitions and results from [45]. For a continuous function $y : \mathbb{R}^d \times [0, T] \rightarrow \mathbb{R}^d$, a general multiple Wiener integral over the subinterval $[t, t+h] \subset [0, T]$ is given by

$$I_{r_1 r_2 \dots r_j}^{t, t+h}(y) = \int_t^{t+h} \int_t^{s_1} \dots \int_t^{s_{j-1}} y(\mathbf{x}_{s_j}, s_j) d\mathbf{w}_{r_1}(s_j) \dots d\mathbf{w}_{r_j}(s_1), \quad (62)$$

where $r_i \in \{0, 1, \dots, d\}$ and $dw_0(s) = ds$. Then we have the following lemma.

Lemma B.9 (Bound of Wiener Integral). *For any function $y : \mathbb{R}^d \times [0, T] \rightarrow \mathbb{R}^d$ that satisfies a growth condition in the form $|y(\mathbf{x}, t)| \leq K(1 + |\mathbf{x}|^2)^{\frac{1}{2}}$, for any $\mathbf{x} \in \mathbb{R}^d$, and any $t \in [0, T]$, $h > 0$ such that $t + h \in [0, T]$, we have that*

$$\mathbb{E} \left[I_{r_1 r_2 \dots r_j}^{t, t+h}(y) | \mathcal{F}_t \right] = 0 \quad \text{if } r_i \neq 0 \text{ for some } i \in \{1, \dots, j\}, \quad (63)$$

$$\left\| I_{r_1 r_2 \dots r_j}^{t, t+h}(y) \right\|_{L_2} = \mathcal{O} \left(h^{l_1 + \frac{l_2}{2}} \right), \quad (64)$$

where l_1 is the number of zero indices and l_2 is the number of non-zero indices r_i .

Lemma B.10 (Property of Lagrange interpolation polynomial). *For $s + 1$ points $(t_{i+1}, y_{i+1}), (t_i, y_i), \dots, (t_{i-(s-1)}, y_{i-(s-1)})$, the Lagrange interpolation polynomial is*

$$L(t) = \sum_{k=i-(s-1)}^{i+1} l_k(t) y_k. \quad (65)$$

Then the following $s+1$ equalities hold

$$\begin{aligned} \sum_{k=i-(s-1)}^{i+1} l_k(u) &= 1, \\ \sum_{k=i-(s-1)}^{i+1} l_k(u) \int_{t_{i-(s-1)}}^{t_k} du_2 &= \int_{t_{i-(s-1)}}^u du_2, \\ &\vdots \\ \sum_{k=i-(s-1)}^{i+1} l_k(u) \int_{t_{i-(s-1)}}^{t_k} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_s} du_{s+1} \cdots du_3 du_2 &= \\ &\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_s} du_{s+1} \cdots du_3 du_2 \end{aligned} \quad (66)$$

Proof. For the first equality, consider $y_k \equiv 1$ for $i - (s - 1) \leq k \leq i + 1$. The Lagrange interpolation polynomial for these y_k s is a constant function $L(t) \equiv 1$. We have $L(u) = \sum_{k=i-(s-1)}^{i+1} l_k(u) = 1$.

For the second equality, consider $y_k = \int_{t_{i-(s-1)}}^{t_k} du_2$. The Lagrange interpolation polynomial for these y_k s is a polynomial of degree 1 $L(t) = t - t_{i-(s-1)}$. We have $L(u) = \sum_{k=i-(s-1)}^{i+1} l_k(u) \int_{t_{i-(s-1)}}^{t_k} du_2 = u - t_{i-(s-1)} = \int_{t_{i-(s-1)}}^u du_2$.

For equalities from the third to the last, without loss of generality, we prove the $p - th$ equality, where $3 \leq p \leq s + 1$. Consider $y_k = \int_{t_{i-(s-1)}}^{t_k} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{p-1}} du_p \cdots du_3 du_2$. The Lagrange interpolation polynomial for these y_k s is a polynomial of degree $p - 1$ $L(t) = \int_{t_{i-(s-1)}}^t \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{p-1}} du_p \cdots du_3 du_2$. We have $L(u) = \int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{p-1}} du_p \cdots du_3 du_2 = \int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{p-1}} du_p \cdots du_3 du_2$. \square

B.4 Proof of Lemma B.7 (for Theorem. 5.1) and Lemma B.8 (for Theorem. 5.2)

To simplify the notation, we will introduce two operators which will appear in the Itô formula. Suppose we have an Itô SDE $d\mathbf{x}_t = f(\mathbf{x}_t, t)dt + g(\mathbf{x}_t, t)d\mathbf{w}_t$ and $h(\mathbf{x}, t)$ is a twice continuously differentiable function. Let $\Gamma_0(\cdot) = \partial_t(\cdot) + \nabla_{\mathbf{x}}(\cdot)f$, $\Gamma_1(\cdot) = \frac{g^2}{2}\Delta(\cdot)$ and $\Gamma_2(\cdot) = \nabla_{\mathbf{x}}(\cdot)g$ in which

$\nabla_{\mathbf{x}}$ is the Jacobian matrix, and Δ is the Laplacian operator. With the notation here, we can express the Itô formula for $h(\mathbf{x}, t)$ as

$$h(\mathbf{x}_t, t) = h(\mathbf{x}_s, s) + \int_s^t (\Gamma_0(h) + \Gamma_1(h)) dt + \int_s^t \Gamma_2(h) d\bar{\mathbf{w}}_t. \quad (67)$$

Given the above lemmas, we will analyze the local error L_{i+1} step by step. Inspired by Theorem B.6, for data-prediction reparameterization model, L_{i+i} can be estimated by decomposing the terms step by step. The first step of decomposition is summarized as the following lemma:

Lemma B.11 (First step of estimating local error L_{i+1} in data-prediction reparameterization model). *Given the exact solution of data prediction model*

$$\begin{aligned} \mathbf{x}_t &= \frac{\sigma_t}{\sigma_s} e^{-\int_{\lambda_s}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \mathbf{x}_s + \sigma_t \mathbf{F}_{\theta}(s, t) + \sigma_t \mathbf{G}(s, t), \\ \mathbf{F}_{\theta}(s, t) &= \int_{\lambda_s}^{\lambda_t} e^{-\int_{\lambda_s}^{\lambda} \tau^2(\bar{\lambda}) d\bar{\lambda}} (1 + \tau^2(\lambda)) e^{\lambda} \mathbf{x}_{\theta}(\mathbf{x}_{\lambda}, \lambda) d\lambda \\ \mathbf{G}(s, t) &= \int_s^t e^{-\int_{\lambda_u}^{\lambda_t} \tau^2(\bar{\lambda}) d\bar{\lambda}} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u, \end{aligned} \quad (68)$$

With proper $b_k, k \in [i - (s - 1), i + 1]$, The local error L_{i+1} in Eq. (56) is

$$L_{i+1} = R_{i+1}^{(1)} + S_{i+1}^{(1)} \quad (69)$$

where

$$\begin{aligned} S_{i+1}^{(1)} &= \mathcal{O} \left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}} \right) \\ R_{i+1}^{(1)} &= \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \\ &\quad \left(\int_{t_k}^{t_{k+1}} \Gamma_0(\mathbf{x}_{\theta}) dt \right) \\ &\quad + \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u \Gamma_0(\mathbf{x}_{\theta}) dt \right) \frac{d\lambda_u}{du} du \\ &\quad - \sum_{j=-1}^{s-1} b_{i-j} \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_0(\mathbf{x}_{\theta}) dt. \end{aligned} \quad (70)$$

Proof. The difference between Eq. (55) and Eq. (56) is that \mathbf{x}_j is our numerical approximation, while \mathbf{x}_{t_j} is the exact solution of SDE Eq. (54) at time $t = t_j$. Substitute the exact solution Eq. (31) of $\mathbf{x}_{t_{i+1}}$, we have

$$\begin{aligned} L_{i+1} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} \mathbf{x}_{t_i} + \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{t_{i+1}} \tau^2(\lambda) d\lambda} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \\ &\quad + \sigma_{t_{i+1}} \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(\lambda_u)) e^{\lambda_u} \mathbf{x}_{\theta}(\mathbf{x}_{\lambda_u}, \lambda_u) d\lambda_u \\ &\quad - \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{t_{i+1}} \tau^2(\lambda) d\lambda} \tau(u) \sqrt{-2 \frac{d\lambda_u}{du}} d\bar{\mathbf{w}}_u \\ &\quad - \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\lambda_u) d\lambda_u} \mathbf{x}_{t_i} - \sum_{j=-1}^{s-1} b_{i-j} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i-j}}, t_{i-j}) \\ &= \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \mathbf{x}_{\theta}(\mathbf{x}_u, u) \frac{d\lambda_u}{du} du \\ &\quad - \sum_{j=-1}^{s-1} b_{i-j} \mathbf{x}_{\theta}(\mathbf{x}_{t_{i-j}}, t_{i-j}). \end{aligned} \quad (71)$$

Let $f(\mathbf{x}, t) = \left(\frac{d \log \alpha_t}{dt} - (1 + \tau^2(t)) \frac{d\lambda_t}{dt} \right) \mathbf{x} + (1 + \tau^2(t)) \alpha_t \mathbf{x}_\theta(\mathbf{x}, t) \frac{d\lambda_t}{dt}$, $g(t) = \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}}$. By Itô's formula [38], we have

$$\begin{aligned} \mathbf{x}_\theta(\mathbf{x}_u, u) &= \mathbf{x}_\theta(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) + \sum_{k=i-(s-1)}^{i-1} \int_{t_k}^{t_{k+1}} (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt \\ &+ \int_{t_i}^u (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt + \sum_{k=i-(s-1)}^{i-1} \int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t + \int_{t_i}^u \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t, \end{aligned} \quad (72)$$

$$\begin{aligned} \mathbf{x}_\theta(\mathbf{x}_{t_{i-j}}, t_{i-j}) &= \mathbf{x}_\theta(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) + \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt \\ &+ \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t, \end{aligned} \quad (73)$$

Substituting Eq. (72) and Eq. (73) into Eq. (71), we have

$$\begin{aligned} &L_{i+1} \\ &= \left(\sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du - \sum_{j=-1}^{s-1} b_{i-j} \right) \mathbf{x}_\theta(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\ &+ \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \\ &\quad \left(\int_{t_k}^{t_{k+1}} (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt \right) \\ &+ \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \left(\int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t \right) \\ &+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt \right) \frac{d\lambda_u}{du} du \\ &+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t \right) \frac{d\lambda_u}{du} du \\ &- \sum_{j=-1}^{s-1} b_{i-j} \left(\sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} (\Gamma_0(\mathbf{x}_\theta) + \Gamma_1(\mathbf{x}_\theta)) dt + \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t \right) \end{aligned} \quad (74)$$

We will divide the local error L_{i+1} into distinct terms. The first term has a coefficient

$$\sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du - \sum_{j=-1}^{s-1} b_{i-j}. \quad (75)$$

By Lemma B.10, b_k constructed by the integral of Lagrange polynomial in Eq. (58) and Eq. (60) satisfies $b_k = \mathcal{O}(h)$ and the coefficient (75) is zero. Furthermore, we have $g(t) = \tau(t) \sigma_t \sqrt{-2 \frac{d\lambda_t}{dt}} =$

$\mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t)\right)$. By Lemma B.9, we have the following estimations

$$\begin{aligned}
& \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \left(\int_{t_k}^{t_{k+1}} \Gamma_1(\mathbf{x}_\theta) dt \right) \\
&= \mathcal{O}\left(\max_{0 \leq t \leq T} \tau^2(t) h^2\right), \\
& \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \left(\int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t \right) \\
&= \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right), \\
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u \Gamma_1(\mathbf{x}_\theta) dt \right) \frac{d\lambda_u}{du} du = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau^2(t) h^2\right), \\
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t \right) \frac{d\lambda_u}{du} du = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right), \\
& \sum_{j=-1}^{s-1} b_{i-j} \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_1(\mathbf{x}_\theta) dt = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau^2(t) h^2\right), \\
& \sum_{j=-1}^{s-1} b_{i-j} \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_2(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right),
\end{aligned} \tag{76}$$

and the summation of the above terms is $S_{i+1}^{(1)} = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right)$.

The remaining terms of local error are

$$\begin{aligned}
R_{i+1}^{(1)} &= \sum_{k=i-(s-1)}^{i-1} \sigma_{t_{i+1}} \left(\int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \frac{d\lambda_u}{du} du \right) \times \left(\int_{t_k}^{t_{k+1}} \Gamma_0(\mathbf{x}_\theta) dt \right) \\
&+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_i}^u \Gamma_0(\mathbf{x}_\theta) dt \right) \frac{d\lambda_u}{du} du \\
&- \sum_{j=-1}^{s-1} b_{i-j} \sum_{k=i-(s-1)}^{i-j-1} \int_{t_k}^{t_{k+1}} \Gamma_0(\mathbf{x}_\theta) dt,
\end{aligned} \tag{77}$$

which completes the proof. \square

The remaining problem is to estimate the $R_{i+1}^{(1)}$ in Eq. (70). We can further expand the term $\Gamma_0(\mathbf{x}_\theta)$ as following

$$\begin{aligned}
& \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_t, t) \\
&= \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) + \int_{t_{i-(s-1)}}^t (\Gamma_0 \Gamma_0(\mathbf{x}_\theta) + \Gamma_1 \Gamma_0(\mathbf{x}_\theta)) dt + \int_{t_{i-(s-1)}}^t \Gamma_2 \Gamma_0(\mathbf{x}_\theta) d\bar{\mathbf{w}}_t.
\end{aligned} \tag{78}$$

Substituting the expansion of $\Gamma_0(\mathbf{x}_\theta)$, we perform the approximation of \tilde{L}_{i+1} , which is summarized with the following lemma:

Lemma B.12 (Second step of estimating L_{i+1} in data-prediction reparameterization model). $R_{i+1}^{(1)}$ in Eq. (70) can be decomposed as

$$R_{i+1}^{(1)} = R_{i+1}^{(2)} + S_{i+1}^{(2)}, \tag{79}$$

where

$$\begin{aligned}
S_{i+1}^{(2)} &= \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{5}{2}}\right) \\
R_{i+1}^{(2)} &= \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u du_2 \right) \frac{d\lambda_u}{du} du \cdot \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
&+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \Gamma_0 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} du_2 \times \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \Gamma_0 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2
\end{aligned} \tag{80}$$

Proof. We start with decomposing the term $R_{i+1}^{(1)}$

$$\begin{aligned}
R_{i+1}^{(1)} &= \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_2}, u_2) du_2 \right) \frac{d\lambda_u}{du} du \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_2}, u_2) du_2 \\
&= \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u du_2 \right) \frac{d\lambda_u}{du} du \cdot \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
&+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\
&\quad \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} (\Gamma_0 \Gamma_0(\mathbf{x}_\theta) + \Gamma_1 \Gamma_0(\mathbf{x}_\theta))(\mathbf{x}_{u_3}, u_3) du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
&+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\
&\quad \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \Gamma_2 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) d\bar{\mathbf{w}}_{u_3} du_2 \right) \frac{d\lambda_u}{du} du \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} du_2 \times \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} (\Gamma_0 \Gamma_0(\mathbf{x}_\theta) + \Gamma_1 \Gamma_0(\mathbf{x}_\theta))(\mathbf{x}_{u_3}, u_3) du_3 du_2 \\
&- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \Gamma_2 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) d\bar{\mathbf{w}}_{u_3} du_2.
\end{aligned} \tag{81}$$

We further estimate the terms with Γ_1 and Γ_2 .

$$\begin{aligned}
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\
& \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \Gamma_1 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
& = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau^2(t) h^3 \right), \\
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\
& \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \Gamma_2 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) d\bar{w}_{u_3} du_2 \right) \frac{d\lambda_u}{du} du \\
& = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{5}{2}} \right), \\
& \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \Gamma_1 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2 = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau^2(t) h^3 \right), \\
& \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \Gamma_2 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) d\bar{w}_{u_3} du_2 = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{5}{2}} \right).
\end{aligned} \tag{82}$$

The summation of the above terms is $S^{(2)} = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{5}{2}} \right)$. Compared with $S^{(1)}$, this term can be omitted.

The remaining local error is

$$\begin{aligned}
R_{i+1}^{(2)} & = \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u du_2 \right) \frac{d\lambda_u}{du} du \\
& \quad \times \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
& \quad + \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\
& \quad \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \Gamma_0 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
& \quad - \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} du_2 \times \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
& \quad - \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \Gamma_0 \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_3}, u_3) du_3 du_2
\end{aligned} \tag{83}$$

which completes the proof. \square

Remark 3. With Lemma B.11 and B.12, the local error L_{i+1} can be decomposed to the term $S_{i+1} = S_{i+1}^{(1)} + S_{i+1}^{(2)}$ and the term $R_{i+1}^{(2)}$. It is clear that $S_{i+1} = \mathcal{O} \left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}} \right)$, and we will show that given b_{i-j} constructed by integral of Lagrange polynomial in Eq. (58) and Eq. (60), $R_{i+1}^{(2)} = \mathcal{O}(h^3)$.

By Lemma B.10, b_k constructed by the integral of Lagrange polynomial in Eq. (58) and Eq. (60) satisfies that the coefficient for $\Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)})$

$$\sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u du_2 \right) \frac{d\lambda_u}{du} du - \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} du_2, \quad (84)$$

equals zero. And the remaining term in $R_{i+1}^{(2)}$ is $\mathcal{O}(h^3)$.

Remark 4. We will show that the local error can be further decomposed such that $L_{i+1} = R_{i+1}^{(s)} + \sum_{j=1}^s S_{i+1}^{(j)}$. In this case $S_{i+1} = \sum_{j=1}^s S_{i+1}^{(j)}$ is the term such that $S_{i+1} = \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right)$, and we will show that by our constructed b_{i-j} , $R_{i+1}^{(s)} = \mathcal{O}(h^{s+1})$.

Lemma B.13 (j -th step of estimating L_{i+1} in data-prediction reparameterization model). For $j \leq s+1$, $R_{i+1}^{(j-1)}$ in Eq. (70) can be decomposed as

$$R_{i+1}^{(j-1)} = R_{i+1}^{(j)} + S_{i+1}^{(j)}, \quad (85)$$

where

$$\begin{aligned} S_{i+1}^{(j)} &= \mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{2j+1}{2}}\right) \\ R_{i+1}^{(j)} &= \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{j-1}} du_j \cdots du_3 du_2 \right) \frac{d\lambda_u}{du} du \\ &\quad \cdot \overbrace{\Gamma_0 \cdots \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)})}^{j-1} \\ &+ \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \\ &\quad \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_j} \overbrace{\Gamma_0 \cdots \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_{j+1}}, u_{j+1})}^j du_{j+1} \cdots du_3 du_2 \right) \frac{d\lambda_u}{du} du \\ &- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{j-1}} du_j \cdots du_3 du_2 \cdot \overbrace{\Gamma_0 \cdots \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)})}^{j-1} \\ &- \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_j} \overbrace{\Gamma_0 \cdots \Gamma_0(\mathbf{x}_\theta)(\mathbf{x}_{u_{j+1}}, u_{j+1})}^j du_{j+1} \cdots du_3 du_2 \end{aligned} \quad (86)$$

Furthermore, given that b_k is constructed by the integral of Lagrange polynomial in Eq. (58) and Eq. (60), $R_{i+1}^{(j)} = \mathcal{O}(h^{j+1})$

Sketch of the proof (1) Use the Itô formula Eq. (67) to expand $\overbrace{\Gamma_0 \cdots \Gamma_0(\mathbf{x}_\theta)}^{j-1}$. (2) Use Lemma B.9 to estimate the stochastic term $S^{(j)}$. For the remaining term $R^{(j)}$, by Lemma B.10, b_k constructed by the integral of Lagrange polynomial in Eq. (58) and Eq. (60) satisfies that the coefficient before

$$\begin{aligned}
& \overbrace{\Gamma_0 \cdots \Gamma_0}^{j-1}(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)}) \\
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{j-1}} du_j \cdots du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
& - \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_{j-1}} du_j \cdots du_3 du_2.
\end{aligned} \tag{87}$$

equals zero. And the remaining term in $R_{i+1}^{(j)}$ is $\mathcal{O}(h^{j+1})$.

The process can be repeated until the coefficient before $\overbrace{\Gamma_0 \cdots \Gamma_0}^s(\mathbf{x}_\theta)(\mathbf{x}_{t_{i-(s-1)}}, t_{i-(s-1)})$ is

$$\begin{aligned}
& \sigma_{t_{i+1}} \int_{t_i}^{t_{i+1}} e^{-\int_{\lambda_u}^{\lambda_{t_{i+1}}} \tau^2(\lambda) d\lambda} (1 + \tau^2(u)) e^{\lambda_u} \left(\int_{t_{i-(s-1)}}^u \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_s} du_{s+1} \cdots du_3 du_2 \right) \frac{d\lambda_u}{du} du \\
& - \sum_{j=-1}^{s-1} b_{i-j} \int_{t_{i-(s-1)}}^{t_{i-j}} \int_{t_{i-(s-1)}}^{u_2} \cdots \int_{t_{i-(s-1)}}^{u_s} du_{s+1} \cdots du_3 du_2.
\end{aligned} \tag{88}$$

which equals zero. And the remaining term R_{i+1}^{s+1} is $\mathcal{O}(h^{s+2})$.

We conclude with the proof of Lemma B.7 and B.8.

Proof for Lemma B.8 (Convergence for s -step SA-Corrector) The stochastic term S_{i+1} can be estimated as $\mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right)$. Lemma B.10 prove that with b_{i-j} defined in Theorem 5.2, the coefficients of Eq. (75), Eq. (84), Eq. (87) and Eq. (88) equal zero. Thus the deterministic term R_{i+1} can be estimated as $\mathcal{O}(h^{s+2})$. The proof is completed.

Proof for Lemma B.7 (Convergence for s -step SA-Predictor) The stochastic term S_{i+1} can be estimated as $\mathcal{O}\left(\max_{0 \leq t \leq T} \tau(t) h^{\frac{3}{2}}\right)$ from Eq. (76) and Eq. (82). Lemma B.10 prove that with b_{i-j} defined in Theorem 5.1, the coefficients of Eq. (75), Eq. (84), Eq. (87) and Eq. (88) equal zero except for the last term. This is because in s -step SA-Predictor we only have s points in contrast to $s+1$ points in s -step SA-Corrector, for which we can only obtain the first s equalities in Lemma B.10. Thus the deterministic term R_{i+1} can be estimated as $\mathcal{O}(h^{s+1})$. The proof is completed.

B.5 Relationship with Existing Samplers

B.5.1 Relationship with DDIM

DDIM [22] generates samples through the following process:

$$\mathbf{x}_{t_{i+1}} = \alpha_{t_{i+1}} \left(\frac{\mathbf{x}_{t_i} - \sigma_{t_i} \boldsymbol{\epsilon}_\theta(\mathbf{x}_{t_i}, t_i)}{\alpha_{t_i}} \right) + \sqrt{1 - \alpha_{t_{i+1}}^2 - \hat{\sigma}_{t_i}^2} \boldsymbol{\epsilon}_\theta(\mathbf{x}_{t_i}, t_i) + \hat{\sigma}_{t_i} \boldsymbol{\xi}, \tag{89}$$

where $\boldsymbol{\xi} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, $\hat{\sigma}_{t_i}$ is a variable parameter. In practice, DDIM introduces a parameter η such that when $\eta = 0$, the sampling process becomes deterministic and when $\eta = 1$, the sampling process coincides with original DDPM [2]. Specifically, $\hat{\sigma}_{t_i} = \eta \sqrt{\frac{1 - \alpha_{t_{i+1}}^2}{1 - \alpha_{t_i}^2}} \left(1 - \frac{\alpha_{t_i}^2}{\alpha_{t_{i+1}}^2} \right)$.

Corollary 5.3 For any η in DDIM, there exists a $\tau_\eta(t) : \mathbb{R} \rightarrow \mathbb{R}$ which is a piecewise constant function such that DDIM- η coincides with our 1-step SA-Predictor when $\tau(t) = \tau_\eta(t)$ with data parameterization of our variance-controlled diffusion SDE.

Proof. Our 1-step *SA-Predictor* when $\tau(t) = \tau, t \in [t_i, t_{i+1}]$ with data parameterization of our variance-controlled diffusion SDE is

$$\begin{aligned} \mathbf{x}_{t_{i+1}} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})} \mathbf{x}_{t_i} + \alpha_{t_{i+1}} \left(1 - e^{-(1+\tau^2)(\lambda_{t_{i+1}} - \lambda_{t_i})}\right) \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) \\ &\quad + \sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}} \boldsymbol{\xi}. \end{aligned} \quad (90)$$

DDIM- η generates samples through the following process

$$\mathbf{x}_{t_{i+1}} = \alpha_{t_{i+1}} \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) + \sqrt{1 - \alpha_{t_{i+1}}^2 - \hat{\sigma}_{t_i}^2} \boldsymbol{\epsilon}_\theta(\mathbf{x}_{t_i}, t_i) + \hat{\sigma}_{t_i} \boldsymbol{\xi}, \quad \hat{\sigma}_{t_i} = \eta \sqrt{\frac{1 - \alpha_{t_{i+1}}^2}{1 - \alpha_{t_i}^2} \left(1 - \frac{\alpha_{t_i}^2}{\alpha_{t_{i+1}}^2}\right)}. \quad (91)$$

If we substitute $\hat{\sigma}_{t_i}$ with $\sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}}$, we can verify that $\sqrt{1 - \alpha_{t_{i+1}}^2 - \hat{\sigma}_{t_i}^2} = \sigma_{t_{i+1}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}$. The DDIM- η then becomes

$$\begin{aligned} \mathbf{x}_{t_{i+1}} &= \alpha_{t_{i+1}} \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) + \sigma_{t_{i+1}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})} \left(\frac{\mathbf{x}_{t_i} - \alpha_{t_i} \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i)}{\sigma_{t_i}}\right) \\ &\quad + \sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}} \boldsymbol{\xi} \\ &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})} \mathbf{x}_{t_i} + \left(\alpha_{t_{i+1}} - \frac{\alpha_{t_i}}{\sigma_{t_i}} \sigma_{t_{i+1}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}\right) \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) \\ &\quad + \sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}} \boldsymbol{\xi} \\ &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})} \mathbf{x}_{t_i} + \alpha_{t_{i+1}} \left(1 - e^{-(1+\tau^2)(\lambda_{t_{i+1}} - \lambda_{t_i})}\right) \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) \\ &\quad + \sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau^2(\lambda_{t_{i+1}} - \lambda_{t_i})}} \boldsymbol{\xi}, \end{aligned} \quad (92)$$

which is exactly the same with our 1-step *SA-Predictor*. To find the τ_η , we solve the relationship

$$\eta \sqrt{\frac{1 - \alpha_{t_{i+1}}^2}{1 - \alpha_{t_i}^2} \left(1 - \frac{\alpha_{t_i}^2}{\alpha_{t_{i+1}}^2}\right)} = \sigma_{t_{i+1}} \sqrt{1 - e^{-2\tau_\eta^2(\lambda_{t_{i+1}} - \lambda_{t_i})}}. \quad (93)$$

The relationship between τ and η is

$$\eta = \sigma_{t_i} \sqrt{\frac{1 - e^{-2\tau_\eta^2(\lambda_{t_{i+1}} - \lambda_{t_i})}}{1 - \frac{\alpha_{t_i}^2}{\alpha_{t_{i+1}}^2}}}, \quad \tau_\eta = \sqrt{\frac{\log\left(1 - \frac{\eta^2}{\sigma_{t_i}^2} \left(1 - \frac{\alpha_{t_i}^2}{\alpha_{t_{i+1}}^2}\right)\right)}{-2(\lambda_{t_{i+1}} - \lambda_{t_i})}}. \quad (94)$$

□

In a concurrent paper [31], Lu *et al.* prove the result that their SDE-DPM-Solver+++ coincides with DDIM with a special η . Their result is a special case of Corollary 5.3 when $\tau_\eta \equiv 1$ and η take a special value, while our result holds for arbitrary η .

B.5.2 Relationship with DPM-Solver++(2M)

DPM-Solver++ [31] is a high-order solver which solves diffusion ODEs for guided sampling. DPM-Solver++(2M) is equivalent to the 2-step Adams-Bashforth scheme combined with the exponential integrator. While our 2-step *SA-Predictor* is also equivalent to the 2-step Adams-Bashforth scheme combined with the exponential integrator when $\tau(t) \equiv 0$. Thus DPM-Solver++(2M) is a special case of our 2-step *SA-Predictor* when $\tau(t) \equiv 0$.

B.5.3 Relationship with UniPC

UniPC [24] is a unified predictor-corrector framework for solving diffusion ODEs. Specifically, UniPC-p uses a p-step Adams-Bashforth scheme combined with the exponential integrator as a predictor and a p-step Adams-Moulton scheme combined with the exponential integrator as a corrector. While our p-step *SA-Predictor* is also equivalent to the p-step Adams-Bashforth scheme combined with the exponential integrator when $\tau(t) \equiv 0$ and our p-step *SA-Corrector* is also equivalent to the p-step Adams-Moulton scheme combined with the exponential integrator when $\tau(t) \equiv 0$. Thus UniPC-p is a special case of our *SA-Solver* when $\tau(t) \equiv 0$ with predictor step p , corrector step p in Algorithm 1.

C Selection on the Magnitude of Stochasticity

In this section, we will show that we choose $\tau(t) \equiv 1$ in a number of NFEs. We will show that under certain conditions, the upper bound of KL divergence between the marginal distribution and the true distribution can be minimized when $\tau(t) \equiv 1$.

Let $p_t(\mathbf{x})$ denotes the marginal distribution of \mathbf{x}_t , by Proposition 4.1, we know that for any bounded measurable function $\tau(t) : [0, T] \rightarrow \mathbb{R}$, the following Reverse SDEs

$$d\mathbf{x}_t = \left[f(t)\mathbf{x}_t - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t) \nabla_{\mathbf{x}} \log p_t(\mathbf{x}_t) \right] dt + \tau(t)g(t)d\bar{\mathbf{w}}_t, \quad \mathbf{x}_T \sim p_T(\mathbf{x}_T), \quad (95)$$

have the same marginal probability distributions. In practice, we substitute $\nabla_{\mathbf{x}} \log p_t(\mathbf{x}_t)$ with $\mathbf{s}_{\theta}(\mathbf{x}_t, t)$ and substitute $p_T(\mathbf{x}_T)$ with π to sample the reverse SDE.

$$d\mathbf{x}_t^{\theta} = \left[f(t)\mathbf{x}_t^{\theta} - \left(\frac{1 + \tau^2(t)}{2} \right) g^2(t) \mathbf{s}_{\theta}(\mathbf{x}_t^{\theta}, t) \right] dt + \tau(t)g(t)d\bar{\mathbf{w}}_t, \quad \mathbf{x}_T^{\theta} \sim \pi, \quad (96)$$

where π is a known distribution, specifically here a Gaussian. We have the following theorem under the Assumption in Appendix A in [3].

Theorem C.1. *Let $p = p_0$ be the data distribution, which is the distribution if we sample from the ground truth reverse SDE (54) at time 0. Let $p_{\theta}^{\tau(t)}$ be the distribution if we sample from the practical reverse SDE (96) at time 0. Under the assumptions above, we have*

$$\begin{aligned} & D_{KL} \left(p \| p_{\theta}^{\tau(t)} \right) \\ & \leq D_{KL} (p_T \| \pi) + \frac{1}{8} \int_0^T \mathbb{E}_{p_t(\mathbf{x})} \left[\left(\tau(t) + \frac{1}{\tau(t)} \right)^2 g^2(t) \|\nabla_{\mathbf{x}} \log p_t(\mathbf{x}) - \mathbf{s}_{\theta}(\mathbf{x}, t)\|^2 \right] dt. \end{aligned} \quad (97)$$

This evidence lower bound (ELBO) is minimized when $\tau(t) \equiv 1$.

Proof. Denote the path measure of Eq. (95) and Eq. (96) as μ and ν respectively. Both μ and ν are uniquely determined by the corresponding SDEs due to assumptions. Consider a Markov kernel $K \left(\{\mathbf{z}_t\}_{t \in [0, T]}, \mathbf{y} \right) = \delta(\mathbf{z}_0 = \mathbf{y})$. Thus we have the following result

$$\int K \left(\{\mathbf{x}_t\}_{t \in [0, T]}, \mathbf{x} \right) d\mu \left(\{\mathbf{x}_t\}_{t \in [0, T]} \right) = p_0(\mathbf{x}), \quad (98)$$

$$\int K \left(\{\mathbf{x}_t^{\theta}\}_{t \in [0, T]}, \mathbf{x} \right) d\nu \left(\{\mathbf{x}_t^{\theta}\}_{t \in [0, T]} \right) = p_{\theta}^{\tau(t)}(\mathbf{x}). \quad (99)$$

By data processing inequality for KL divergence

$$\begin{aligned} & D_{KL} \left(p \| p_{\theta}^{\tau(t)} \right) = D_{KL} \left(p_0 \| p_{\theta}^{\tau(t)} \right) \\ & = D_{KL} \left(\int K \left(\{\mathbf{x}_t\}_{t \in [0, T]}, \mathbf{x} \right) d\mu \left(\{\mathbf{x}_t\}_{t \in [0, T]} \right) \parallel \int K \left(\{\mathbf{x}_t^{\theta}\}_{t \in [0, T]}, \mathbf{x} \right) d\nu \left(\{\mathbf{x}_t^{\theta}\}_{t \in [0, T]} \right) \right) \\ & \leq D_{KL} (\mu \| \nu). \end{aligned} \quad (100)$$

By the chain rule of KL divergence, we have

$$D_{KL}(\boldsymbol{\mu} \parallel \boldsymbol{\nu}) = D_{KL}(p_T \parallel \pi) + \mathbb{E}_{\mathbf{z} \sim p_T} [D_{KL}(\boldsymbol{\mu}(\cdot | \mathbf{x}_T = \mathbf{z}) \parallel \boldsymbol{\nu}(\cdot | \mathbf{x}_T^\theta = \mathbf{z}))]. \quad (101)$$

By Girsanov Theorem, $D_{KL}(\boldsymbol{\mu}(\cdot | \mathbf{x}_T = \mathbf{z}) \parallel \boldsymbol{\nu}(\cdot | \mathbf{x}_T^\theta = \mathbf{z}))$ can be computed as

$$\begin{aligned} & D_{KL}(\boldsymbol{\mu}(\cdot | \mathbf{x}_T = \mathbf{z}) \parallel \boldsymbol{\nu}(\cdot | \mathbf{x}_T^\theta = \mathbf{z})) \\ &= \mathbb{E}_\mu \left[\int_0^T \frac{1}{2} \left(\tau(t) + \frac{1}{\tau(t)} \right) g(t) (\nabla_{\mathbf{x}} \log p_t(\mathbf{x}) - \mathbf{s}_\theta(\mathbf{x}, t)) d\bar{\mathbf{w}}_t \right] \\ & \quad + \mathbb{E}_\mu \left[\frac{1}{2} \int_0^T \frac{1}{4} \left(\tau(t) + \frac{1}{\tau(t)} \right)^2 g^2(t) \|\nabla_{\mathbf{x}} \log p_t(\mathbf{x}) - \mathbf{s}_\theta(\mathbf{x}, t)\|^2 dt \right] \\ &= \frac{1}{8} \int_0^T \mathbb{E}_{p_t(\mathbf{x})} \left[\left(\tau(t) + \frac{1}{\tau(t)} \right)^2 g^2(t) \|\nabla_{\mathbf{x}} \log p_t(\mathbf{x}) - \mathbf{s}_\theta(\mathbf{x}, t)\|^2 \right] dt \end{aligned} \quad (102)$$

□

D Implementation Details

For our 2-step SA-Predictor and 1-step SA-Corrector, we find that the coefficient will degenerate to a simple case.

For 2-step SA-Predictor, assume on $[t_i, t_{i+1}]$, $\tau(t) = \tau$ is a constant,

$$b_i = e^{-\lambda_{t_{i+1}} \tau^2} \sigma_{t_{i+1}} (1 + \tau^2) \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{(1+\tau^2)\lambda} \frac{\lambda - \lambda_{t_{i-1}}}{\lambda_{t_i} - \lambda_{t_{i-1}}} d\lambda, \quad (103)$$

$$b_{i-1} = e^{-\lambda_{t_{i+1}} \tau^2} \sigma_{t_{i+1}} (1 + \tau^2) \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{(1+\tau^2)\lambda} \frac{\lambda - \lambda_{t_i}}{\lambda_{t_{i-1}} - \lambda_{t_i}} d\lambda, \quad (104)$$

we have

$$\begin{aligned} \mathbf{x}_{t_{i+1}} &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\bar{\lambda}) d\bar{\lambda}} \mathbf{x}_{t_i} + b_i \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) + b_{i-1} \mathbf{x}_\theta(\mathbf{x}_{t_{i-1}}, t_{i-1}) + \tilde{\sigma}_i \boldsymbol{\xi} \\ &= \frac{\sigma_{t_{i+1}}}{\sigma_{t_i}} e^{-\int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} \tau^2(\bar{\lambda}) d\bar{\lambda}} \mathbf{x}_{t_i} + (b_i + b_{i-1}) \mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) \\ & \quad - b_{i-1} (\mathbf{x}_\theta(\mathbf{x}_{t_i}, t_i) - \mathbf{x}_\theta(\mathbf{x}_{t_{i-1}}, t_{i-1})) + \tilde{\sigma}_i \boldsymbol{\xi}. \end{aligned} \quad (105)$$

Let $h = \lambda_{t_{i+1}} - \lambda_{t_i}$, we have

$$\begin{aligned} b_i + b_{i-1} &= e^{-\lambda_{t_{i+1}} \tau^2} \sigma_{t_{i+1}} (1 + \tau^2) \int_{\lambda_{t_i}}^{\lambda_{t_{i+1}}} e^{(1+\tau^2)\lambda} d\lambda \\ &= \alpha_{t_{i+1}} (1 - e^{-h(1+\tau^2)}) \\ b_{i-1} &= \alpha_{t_{i+1}} \frac{e^{-(1+\tau^2)h} + (1 + \tau^2)h - 1}{(1 + \tau^2)(\lambda_{t_i} - \lambda_{t_{i-1}})} \\ &= \frac{\alpha_{t_{i+1}}}{\lambda_{t_i} - \lambda_{t_{i-1}}} \frac{1 - (1 + \tau^2)h + \frac{1}{2}(1 + \tau^2)^2 h^2 + \mathcal{O}(h^3) + (1 + \tau^2)h - 1}{1 + \tau^2} \\ &= \frac{\alpha_{t_{i+1}}}{\lambda_{t_i} - \lambda_{t_{i-1}}} \frac{1}{2} (1 + \tau^2) h^2 + \mathcal{O}(h^3) \end{aligned} \quad (106)$$

Thus we implement \hat{b}_{i-1} as $\frac{\alpha_{t_{i+1}}}{\lambda_{t_i} - \lambda_{t_{i-1}}} \frac{1}{2} (1 + \tau^2) h^2$ and \hat{b}_i as $\alpha_{t_{i+1}} (1 - e^{-h(1+\tau^2)}) - \hat{b}_{i-1}$. Note that substituting b_i, b_{i-1} as \hat{b}_i, \hat{b}_{i-1} will maintain the convergence order result of 2-step SA-Predictor since the modified term is $\mathcal{O}(h^3)$. The implementation detail for 1-step SA-Corrector is technically the same.

E Experiment Details

E.1 Details on $\tau(t)$, Predictor Steps and Corrector Steps

CIFAR10 32x32 For the CIFAR10 experiment in Section 6.4, we use the pretrained baseline-unconditional-VE model⁴ from [27]. It’s an unconditional model with VE noise schedule. To fairly compare with results in [27], we use a piecewise constant function $\tau(t)$ inspired by [27]. Concretely, denoting $\sigma_t^{EDM} = \frac{\sigma_t}{\alpha_t}$, our $\tau(t)$ is set to be a constant τ in the interval $[(\sigma^{EDM})^{-1}(0.05), (\sigma^{EDM})^{-1}(1)]$ and to be zero outside the interval. We find empirically that this piecewise constant function setting makes our *SA-Solver* converge better, especially in large noise scale cases. We use a 3-step SA-Predictor and a 3-step SA-Corrector. For the CIFAR10 experiment in Section 6.2 and 6.5, we also use the piecewise constant function $\tau(t)$ as above. The predictor steps and corrector steps vary to verify the effectiveness of our proposed method in Section 6.2, while they are both set to be 3-steps in Section 6.5.

ImageNet 64x64 For the ImageNet 64x64 experiment in Section 6.4, we use the pretrained model⁵ from [4]. It’s a conditional model with VP cosine noise schedule. To fairly compare with results in [27], we use a piecewise constant function $\tau(t)$ inspired by [27]. Concretely, denoting $\sigma_t^{EDM} = \frac{\sigma_t}{\alpha_t}$, our $\tau(t)$ is set to be a constant τ in the interval $[(\sigma^{EDM})^{-1}(0.05), (\sigma^{EDM})^{-1}(50)]$ and to be zero outside the interval. We find empirically that this piecewise constant function setting makes our *SA-Solver* converge better, especially in large noise scale cases. We use a 3-step SA-Predictor and a 3-step SA-Corrector.

Other experiments For other experiments, we use a constant function $\tau(t) \equiv \tau$. It’s generally not the optimal choice for each individual task, thus further fine-grained tuning has the potential to improve the results. We aim to report the result of our *SA-Solver* without extra hyperparameter tuning. We use a 3-step SA-Predictor and a 3-step SA-Corrector under 20 NFEs and 2-step SA-Predictor and a 1-step SA-Corrector beyond 20 NFEs.

E.2 Details on Pretrained Models and Settings

CIFAR10 32x32 For the CIFAR10 experiment, we use the pretrained baseline-unconditional-VE model⁶ from [27]. It’s an unconditional model with VE noise schedule. To fairly compare with results in [27], we follow the time step schedule in it. Specifically, we set $\sigma_{min} = 0.02$ and $\sigma_{max} = 80$ and select the step by $\sigma_i = (\sigma_{max}^{\frac{1}{7}} + \frac{i}{N-1}(\sigma_{min}^{\frac{1}{7}} - \sigma_{max}^{\frac{1}{7}}))^7$ for *SA-Solver* and UniPC. We directly report the results of the deterministic sampler and stochastic sampler of EDM. To make it a strong baseline, we report the results of the optimal setting for 4 hyper-parameters $\{S_{churn}, S_{tmin}, S_{tmax}, S_{noise}\}$ and report its lowest observed FID. While for *SA-Solver* and UniPC, we report the averaged observed FID.

ImageNet 64x64 For the ImageNet 64x64 experiment, we use the pretrained model⁷ from [4]. It’s a conditional model with VP cosine noise schedule. To fairly compare with results in [27], we follow the time step schedule in it and use conditional sampling. Specifically, we set $\sigma_{min} = 0.0064$ and $\sigma_{max} = 80$ and select the step by $\sigma_i = (\sigma_{max}^{\frac{1}{7}} + \frac{i}{N-1}(\sigma_{min}^{\frac{1}{7}} - \sigma_{max}^{\frac{1}{7}}))^7$ for *SA-Solver*, UniPC, DPM-Solver and DDIM. We directly report the results of the deterministic sampler and stochastic sampler of EDM. To make it a strong baseline, we report the results of the optimal setting for 4 hyper-parameters $\{S_{churn}, S_{tmin}, S_{tmax}, S_{noise}\}$ and report its lowest observed FID. While for *SA-Solver*, UniPC, DPM-Solver and DDIM, we report the averaged observed FID.

⁴<https://nvlabs-fi-cdn.nvidia.com/edm/pretrained/baseline/baseline-cifar10-32x32-uncond-ve.pkl>

⁵https://openaipublic.blob.core.windows.net/diffusion/jul-2021/64x64_diffusion.pt

⁶<https://nvlabs-fi-cdn.nvidia.com/edm/pretrained/baseline/baseline-cifar10-32x32-uncond-ve.pkl>

⁷https://openaipublic.blob.core.windows.net/diffusion/jul-2021/64x64_diffusion.pt

ImageNet 256x256 For the ImageNet 256x256 experiment, we use three different pretrained models: LDM⁸(VP, handcrafted noise schedule) from [5], DiT-XL/2⁹(VP, linear noise schedule) from [41], Min-SNR¹⁰(VP, cosine noise schedule) from [42]. We use classifier-free guidance of scale $s = 1.5$ and a uniform time step schedule because it’s the most common setting for guided sampling for ImageNet 256x256.

ImageNet 512x512 For the ImageNet 256x256 experiment, we use the pre-trained model: DiT-XL/2¹¹ from [41]. We use classifier-free guidance of scale $s = 1.5$ and a uniform time step schedule following the settings of DiT [41].

LSUN Bedroom 256x256 For the LSUN Bedroom 256x256 experiment, we use the pretrained model¹² from [4]. We use unconditional sampling and a uniform lambda step schedule from [23].

F Additional Results

We include the detailed FID results in Figure 1, Figure 2 and Figure 4 in the tables 4 to 14. The ablation study shows that stochasticity indeed helps improve sample quality. We find that for small NFEs, the magnitude of stochasticity should be small while for large NFEs, large magnitude of stochasticity helps improve sample quality. It can also be observed that in latent space, SDE converges faster as in Table 13. With only 10 NFEs, $\tau = 0.6$ is better than $\tau = 0$. With 20 NFEs, our *SA-Solver* can achieve 3.87 FID, which outperforms all ODE samplers even with far more steps.

Table 4: Sample quality measured by FID \downarrow on CIFAR10 32x32 dataset (VE-baseline model from [27]) varying the number of function evaluations (NFE). For the results from EDM[†], we reported its lowest observed FID.

Method \ NFE	11	15	23	31	47	63	95
DDIM($\eta = 0$)	18.28	12.23	7.93	6.45	5.27	4.83	4.42
DPM-Solver	9.26	5.13	4.52	4.30	4.02	3.97	3.94
UniPC	6.42	5.02	4.19	4.00	3.91	3.90	3.89
EDM(ODE) [†]	13.46	5.62	4.04	3.82	3.79	3.80	3.79
EDM(SDE) [†]	23.94	8.94	4.73	3.95	3.59	3.36	3.06
SA-Solver	6.46	4.91	3.77	3.40	2.92	2.74	2.63

Table 5: Sample quality measured by FID \downarrow on CIFAR10 32x32 dataset (VE-baseline model from [27]) varying the number of function evaluations (NFE) and the magnitude of stochasticity (τ).

SA-Solver \ NFE	11	15	23	31	47	63	95
$\tau = 0.0$	6.46	5.06	4.22	4.02	3.93	3.92	3.91
$\tau = 0.2$	6.54	5.01	4.14	3.95	3.89	3.84	3.83
$\tau = 0.4$	6.79	4.91	4.03	3.81	3.76	3.74	3.67
$\tau = 0.6$	7.34	4.91	3.85	3.65	3.60	3.56	3.57
$\tau = 0.8$	8.61	5.28	3.77	3.48	3.45	3.43	3.50
$\tau = 1.0$	10.89	6.52	3.98	3.40	3.21	3.25	3.29
$\tau = 1.2$	14.49	9.33	5.19	3.69	3.00	3.03	3.07
$\tau = 1.4$	20.19	13.76	7.60	4.91	2.92	2.86	2.93
$\tau = 1.6$	27.90	20.51	11.89	8.07	3.25	2.74	2.80
$\tau = 1.8$	36.26	29.43	18.13	14.00	4.60	2.83	2.63

⁸<https://ommer-lab.com/files/latent-diffusion/nitro/cin/model.ckpt>

⁹<https://dl.fbaipublicfiles.com/DiT/models/DiT-XL-2-256x256.pt>

¹⁰https://github.com/TiankaiHang/Min-SNR-Diffusion-Training/releases/download/v0.0.0/ema_0.9999_xl.pt

¹¹<https://dl.fbaipublicfiles.com/DiT/models/DiT-XL-2-512x512.pt>

¹²https://openaipublic.blob.core.windows.net/diffusion/jul-2021/lsun_bedroom.pt

Table 6: Sample quality measured by FID \downarrow on ImageNet 64x64 dataset (model from [4]) varying the number of function evaluations (NFE). For the results from EDM[†], we reported its lowest observed FID.

Method \ NFE	15	23	31	47	63	95
DDIM($\eta = 0$)	8.48	5.39	4.27	3.46	3.17	2.95
DPM-Solver	3.49	3.04	2.88	2.80	2.76	2.74
UniPC	3.51	2.84	2.75	2.72	2.71	2.72
EDM(ODE) [†]	4.78	3.12	2.84	2.73	2.73	2.67
EDM(SDE) [†]	8.94	4.30	3.40	2.72	2.44	2.22
SA-Solver	3.41	2.61	2.23	1.95	1.88	1.81

Table 7: Sample quality measured by FID \downarrow on ImageNet 64x64 dataset (model from [4]) varying the number of function evaluations (NFE) and the magnitude of stochasticity (τ).

SA-Solver \ NFE	15	23	31	47	63	95
$\tau = 0.0$	3.48	2.72	2.72	2.66	2.64	2.71
$\tau = 0.2$	3.41	2.80	2.63	2.63	2.64	2.60
$\tau = 0.4$	3.52	2.70	2.51	2.51	2.49	2.49
$\tau = 0.6$	3.98	2.61	2.44	2.39	2.34	2.35
$\tau = 0.8$	5.80	2.68	2.32	2.24	2.19	2.21
$\tau = 1.0$	10.06	3.38	2.23	2.09	2.08	2.08
$\tau = 1.2$	18.39	5.52	2.52	1.95	1.97	2.00
$\tau = 1.4$	32.42	10.37	3.83	2.05	1.89	1.89
$\tau = 1.6$	52.31	19.64	7.10	2.60	1.88	1.81

G Additional Samples

We include additional samples in this section. In Figure 5 and Figure 6 we compare samples of our proposed *SA-Solver* with other diffusion samplers. In Figure 7 and Figure 8, we compare samples of our proposed *SA-Solver* under different NFEs and τ . In Figure 9 and Figure 10, we compare samples of our proposed *SA-Solver* with other diffusion samplers on text-to-image tasks. Our *SA-Solver* can generate more diverse samples with more details.

Table 8: Sample quality measured by FID \downarrow on CIFAR10 32x32 dataset (model trained by ourselves; see Section 6.5) varying the sampling method and the training epoch.

method (NFE = 31) \ epoch	1250	1300	1350	1400	1450	1500
DDIM	39.32	29.79	19.59	12.98	8.63	6.64
DPM-Solver	30.57	22.11	13.85	8.85	5.68	4.55
EDM(ODE)	27.51	19.82	12.37	8.03	5.33	4.32
SA-Solver($\tau = 0.6$)	20.55	14.89	9.71	6.55	4.61	4.08
SA-Solver($\tau = 1.0$)	13.62	10.01	6.79	4.81	3.70	3.47

Table 9: Sample quality measured by FID \downarrow on ImageNet 256x256 dataset (model trained by ourselves; see Section 6.5) varying the sampling method and the training epoch.

method (NFE = 40) \ epoch	50	100	150	200	250
DDIM	19.40	9.61	6.75	5.86	5.12
DPM-Solver	18.75	8.96	6.15	5.28	4.62
SA-Solver($\tau = 0.4$)	17.93	8.39	5.69	4.84	4.24
SA-Solver($\tau = 0.8$)	16.57	7.54	5.15	4.48	3.99

Table 10: Sample quality measured by FID \downarrow on ImageNet 256x256 dataset(model from [5]) varying the number of function evaluations (NFE).

Method \ NFE	5	10	20	40	60	80	100
DDIM($\eta = 0$)	58.68	16.32	6.82	4.71	4.45	4.28	4.23
DPM-Solver	166.88	6.19	5.51	4.17	4.18	4.21	4.15
UniPC	12.79	4.96	4.21	4.14	4.12	4.09	4.10
DDIM($\eta = 1$)	138.91	50.05	14.60	6.09	4.56	4.12	3.87
SA-Solver	11.46	4.82	3.88	3.47	3.37	3.37	3.33

Table 11: Ablation study on the effect of the magnitude of stochasticity using *SA-Solver*. Sample quality measured by FID \downarrow on CIFAR10 32x32 dataset(model from [27]) varying the number of function evaluations (NFE) and the magnitude of stochasticity(τ).

τ \ NFE	15	23	31	47	63	95	127
0	4.84	4.11	3.94	3.86	3.88	3.87	3.87
0.2	4.96	4.04	3.84	3.75	3.74	3.79	3.75
0.4	5.27	4.00	3.87	3.64	3.71	3.70	3.62
0.6	6.05	3.95	3.61	3.49	3.46	3.53	3.43
0.8	7.40	4.12	3.53	3.28	3.37	3.32	3.30
1.0	10.00	4.49	3.41	3.18	3.24	3.17	3.15
1.2	13.58	5.14	3.59	3.10	3.02	2.97	3.05
1.4	17.88	6.55	3.94	3.04	3.01	2.89	2.95
1.6	22.42	8.44	4.69	3.20	3.02	2.94	2.89

Table 12: Ablation study on the effect of the magnitude of stochasticity using *SA-Solver*. Sample quality measured by FID \downarrow on ImageNet 64x64 dataset(model from [4]) varying the number of function evaluations (NFE) and the magnitude of stochasticity(τ).

$\tau \setminus$ NFE	20	40	60	80	100
0	3.30	2.83	2.78	2.79	2.82
0.2	3.32	2.77	2.72	2.74	2.79
0.4	3.37	2.68	2.63	2.62	2.59
0.6	3.61	2.57	2.49	2.49	2.47
0.8	4.19	2.51	2.40	2.34	2.30
1.0	5.55	2.54	2.32	2.21	2.20
1.2	7.93	2.77	2.29	2.14	2.14
1.4	11.55	3.20	2.40	2.14	2.08
1.6	16.15	3.97	2.60	2.20	2.09

Table 13: Ablation study on the effect of the magnitude of stochasticity using *SA-Solver*. Sample quality measured by FID \downarrow on ImageNet 256x256 dataset(model from [5]) varying the number of function evaluations (NFE) and the magnitude of stochasticity(τ).

$\tau \setminus$ NFE	5	10	20	40	60	80	100
0	11.46	5.04	4.30	4.16	4.12	4.10	4.16
0.2	11.88	4.89	4.29	4.05	4.02	4.01	4.03
0.4	12.69	4.84	4.14	3.86	3.84	3.83	3.84
0.6	14.84	4.82	3.99	3.63	3.62	3.63	3.61
0.8	18.82	5.09	3.87	3.55	3.50	3.47	3.47
1.0	25.96	6.06	3.88	3.47	3.41	3.39	3.38
1.2	37.20	8.23	3.92	3.47	3.37	3.37	3.33
1.4	53.03	12.93	4.08	3.53	3.40	3.38	3.36
1.6	71.30	24.08	4.43	3.56	3.44	3.45	3.33

Table 14: Ablation study on the effect of the magnitude of stochasticity using *SA-Solver*. Sample quality measured by FID \downarrow on LSUN Bedroom 256x256 dataset(model from [4]) varying the number of function evaluations (NFE) and the magnitude of stochasticity(τ).

$\tau \setminus$ NFE	20	40	60	80	100
0	3.60	3.14	3.06	3.09	3.07
0.2	3.51	3.12	3.00	2.99	2.99
0.4	3.70	3.09	2.97	3.03	3.16
0.6	4.10	3.08	2.95	2.99	3.03
0.8	4.75	3.11	2.97	2.89	2.99
1.0	6.18	3.28	2.98	2.90	2.91
1.2	8.54	3.53	3.12	2.86	3.00
1.4	12.14	4.25	3.24	2.98	2.93
1.6	16.63	5.50	3.75	3.18	3.10

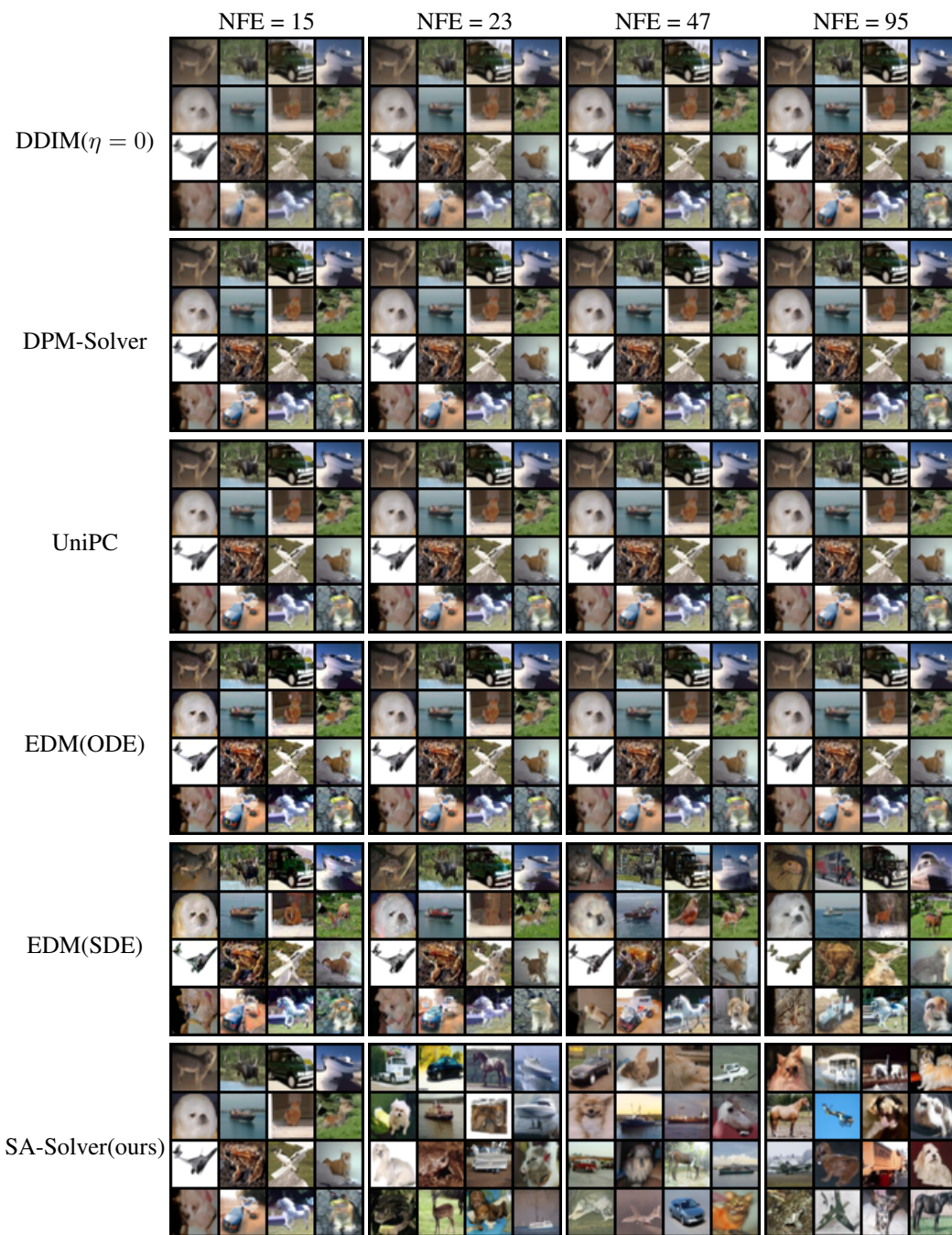


Figure 5: Samples by DDIM, DPM-Solver, UniPC, EDM(ODE), EDM(SDE) and our SA-Solver with 15, 23, 47, 95 NFEs with the same random seed from CIFAR10 32x32 VE baseline model [27]

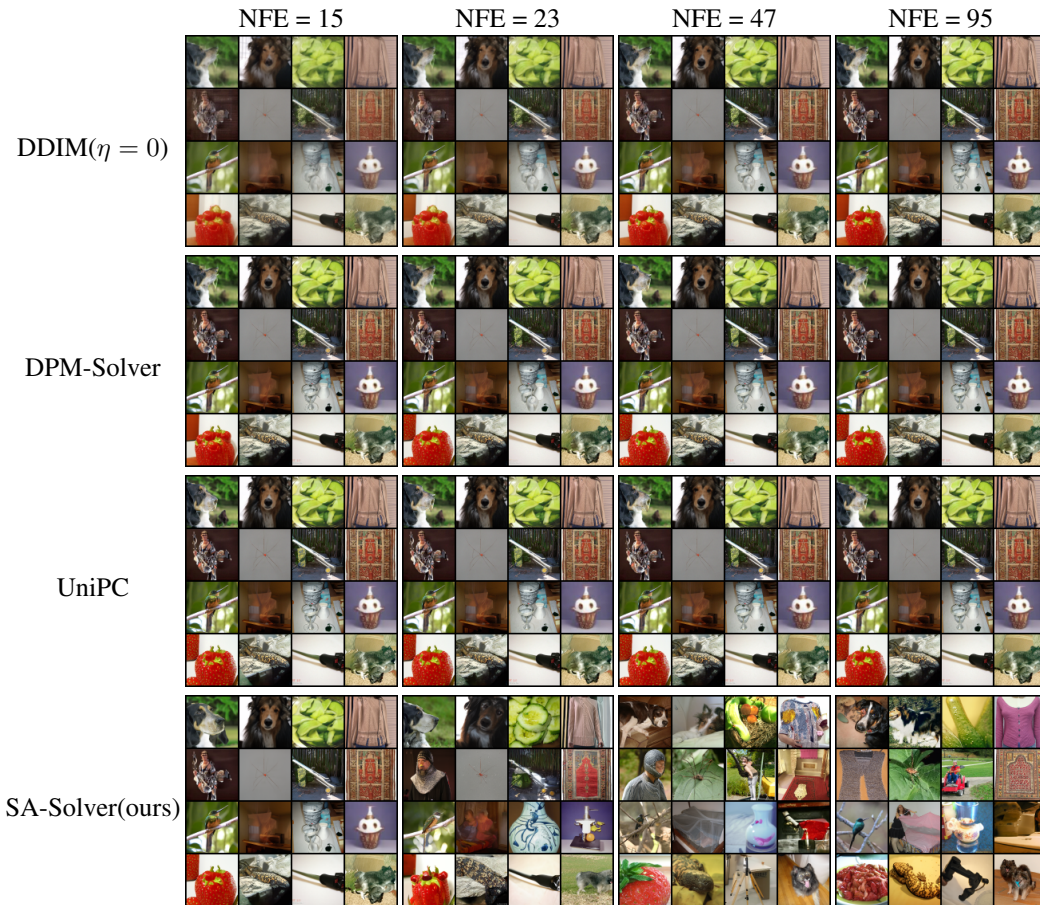


Figure 6: Samples by DDIM, DPM-Solver, UniPC, and our SA-Solver with 15, 23, 47, 95 NFEs with the same random seed from ImageNet 64x64 model [27](conditional sampling)



Figure 7: Samples by SA-Solver with 20, 40, 60, 100 NFEs varying stochasticity(τ) with the same random seed from LSUN-Bedroom 256x256 model [4](unconditional sampling).

NFE = 60



Figure 8: Samples by SA-Solver with 60 NFEs varying stochasticity(τ) with the same random seed from ImageNet 512x512 DiT model [41] with classifier-free guidance scale $s = 4.0$ (default setting to show image).

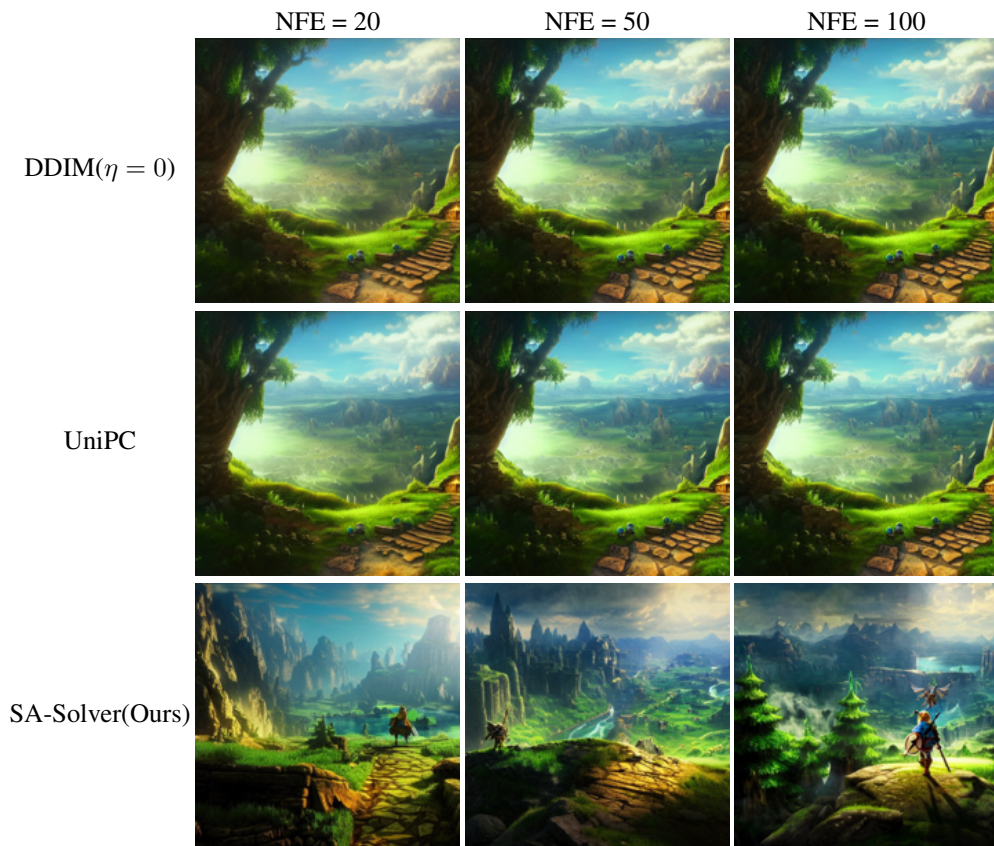


Figure 9: Samples using Stable-Diffusion v1.5 [5] with a classifier-free guidance scale 7.5 with different solvers and NFEs. Prompt: The Legend of Zelda landscape atmospheric, hyper realistic, 8k, epic composition, cinematic, octane render, artstation landscape vista photography by Carr Clifton Galen Rowell, 16K resolution, Landscape veduta photo by Dustin Lefevre tdraw, 8k resolution, detailed landscape painting by Ivan Shishkin, DeviantArt, Flickr, rendered in Enscape, Miyazaki, Nausicaa Ghibli, Breath of The Wild, 4k detailed post processing, artstation, rendering by octane, unreal engine.

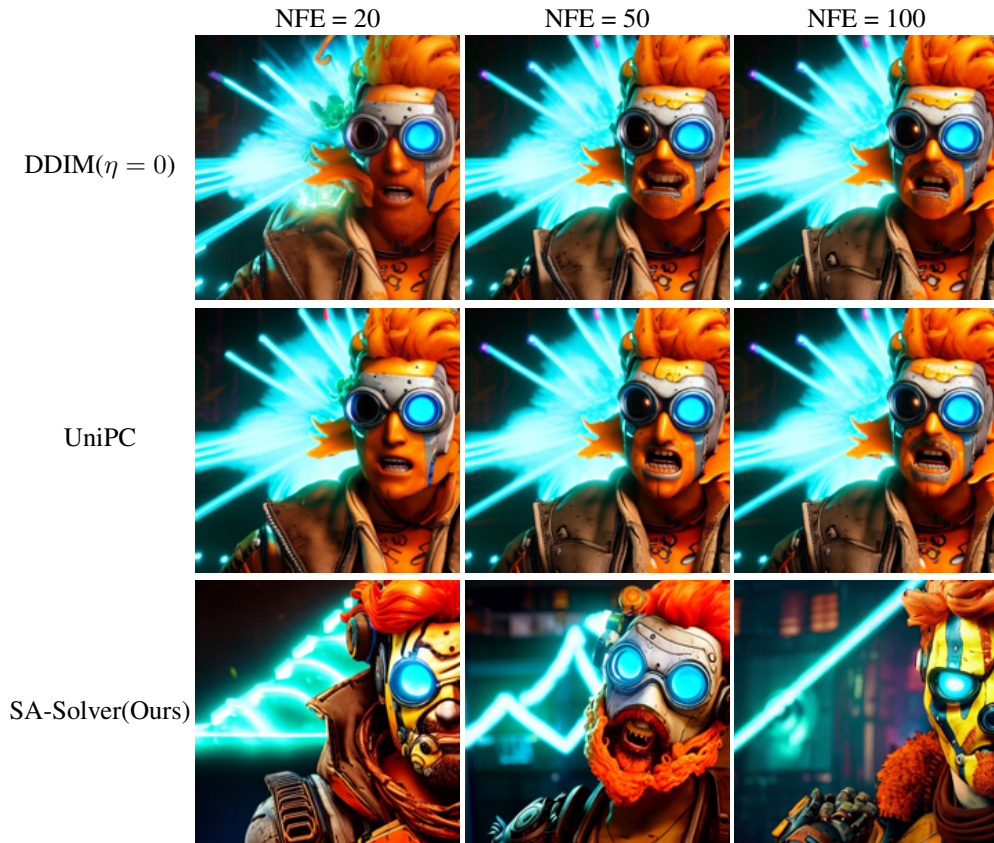


Figure 10: Samples using Stable-Diffusion v1.5 [5] with a classifier-free guidance scale 7.5 with different solvers and NFEs. Prompt: glowwave portrait of curly orange haired mad scientist man from borderlands 3, au naturel, hyper detailed, digital art, trending in artstation, cinematic lighting, studio quality, smooth render, unreal engine 5 rendered, octane rendered, art style by pixar dreamworks warner bros disney riot games and overwatch.