

1 We thank the reviewers and the editor for their reviews and helpful comments, which will improve our manuscript. We
2 are grateful to see Reviewer 2 (R2) write that “this paper is likely to become influential” and “is novel, well-written
3 and important.” We take seriously R2’s suggestion to change our method’s acronym and thank him/her for pointing out
4 several references, which we have added to the paper. We have also followed the reviewer’s suggestion to illustrate
5 CQR using the heavy-tailed Cauchy distribution. As predicted by R2, this experiment indeed shows a clear advantage
6 of CQR over split and locally adaptive conformal prediction. We have added this example to our manuscript.

7 Reviewer 1 (R1) observes that, under mild conditions, classical conformal prediction is valid for any regression
8 algorithm and any conformity score function. This is obviously true and is noted in Section 5 of the manuscript.
9 Therefore, the crucial question is this: **which regression algorithm and conformity score should one then use?** In
10 this respect, our work marks a significant departure from the whole body of research built on the original version of
11 conformal prediction, for *conditional mean regression*. We argue that the new types of conformity scores we develop
12 improve significantly on the state of the art. We cannot put it better than R2: “although it [the use of quantile regression]
13 appears to be a simple modification from hindsight, it is non-trivial from foresight because quantile based methods
14 is highly adaptive to the heteroscedasticity, or more generally distributional heterogeneity, which is ubiquitous in
15 real-world applications.” This is the main point of our paper. Conformal inference is a beautiful idea but what if it had
16 been implemented with subpar tools all along? In particular, why estimate the mean if the goal is to estimate quantiles?

17 To understand the limitations of classical conformal prediction, consider heteroscedastic data with outliers. Suppose we
18 have complete knowledge of the conditional distribution $Y|X$, so that no learning is required. In this idealized setting,
19 the usual conformity score becomes $R_i = |Y_i - \mu(X_i)|$, where μ is the *true* conditional mean regression function.
20 This score would never yield optimal prediction intervals on heteroscedastic data! It is optimal only in the restrictive
21 setting of a location model $Y = \mu(X) + \epsilon$, where the noise ϵ follows a symmetric, unimodal density function [1].
22 As for locally adaptive split conformal prediction, we argue in our paper that scaling residuals by their variance is
23 also suboptimal, even if its limitations are less severe. In the present idealized setting, this method can only construct
24 intervals that are adaptive to the location and local variance in $Y|X$. Distributional heterogeneity does not necessarily
25 arise from such a location-scale family, as illustrated in our synthetic simulation (Figure 1). Furthermore, the locally
26 adaptive score performs poorly on data with outliers, as confirmed by R2, who kindly guided us to design an additional
27 experiment that corroborates this point.

28 In the ideal setting where the conditional distribution is known, *CQR would construct exact prediction intervals*
29 *reflecting the intrinsic predictive uncertainty, while achieving any desired coverage level*. To see this, recall that the
30 endpoints of the CQR prediction interval would be the true lower and upper conditional quantiles, and so the correction
31 term $Q_{1-\alpha}(E, \mathcal{I}_2)$ would be equal to zero. This property stands in contrast to *all previous conformal prediction methods,*
32 *which have generally nonzero correction terms even in the idealized setting*. (We thank R1 for raising concerns about the
33 novelty of our paper and its importance; we have modified our manuscript to include this crucial discussion.) When the
34 conditional distribution is unknown, our method improves accuracy. To quote R2: “[CQR] may influence the machine
35 learning architecture for problems [with] continuous outcome. For instance, the last layer is typically the L2 loss, which
36 gives no easy way to assess decision uncertainty with theoretical guarantee. However, replacing the last layer by pinball
37 loss with two different quantiles would automatically provide an assessment of uncertainty and the conformalization
38 step provides the theoretical guarantee in finite samples even in presence of arbitrary model misspecification.”

39 R1 comments that the length of the prediction intervals constructed by the full conformal method is not fixed, whereas
40 in split conformal it is equal to $2Q_{1-\alpha}(R, \mathcal{I}_2)$. This is true. To quote from [1], however: “for full conformal, the
41 width can vary slightly as X varies, but the difference is often negligible as long as the fitting method is moderately
42 stable.” We added a similar comment to our paper. That being said, the enormous computational cost of full conformal
43 compared to split conformal means that *any comparison between the two is of very limited practical interest*.

44 R1 also wonders why the criticism about the fixed-length interval does not apply to CQR, as its intervals have length
45 *lower bounded* by $2Q_{1-\alpha}(E, \mathcal{I}_2)$, independently of X_{n+1} . Our Figure 1 clearly shows that the intervals constructed by
46 CQR do *not* have fixed length. As to why, notice that the interval length for split conformal is *equal to*, not merely lower
47 bounded by, $2Q_{1-\alpha}(R, \mathcal{I}_2)$, whereas in our case, the quantity $Q_{1-\alpha}(E, \mathcal{I}_2)$ can be positive, zero, or even negative,
48 depending on the calibration of the quantile regression method. When perfectly calibrated, as in the ideal setting, this
49 quantity will be zero, making the lower bound trivial. We appreciate the reviewer’s suggestion to integrate the locally
50 adaptive approach into our framework. However, we are doubtful that doing so would improve its performance, since,
51 as R2 notes, our paper “provides extensive high-quality numerical experiments, which clearly demonstrate the superior
52 performance and adaptivity of conformalized quantile regression.”

53 Reviewer 3 suggests that we construct a loss function formulating CQR as an optimization problem. In fact, we use the
54 pinball loss function to estimate the conditional quantiles, which can be viewed as CQR’s objective function.

55 [1] Lei et al. Distribution-free predictive inference for regression. *JASA*, 113(523):1094–1111, 2018.